## Introduction à la statistique médicale

## Statistical Parametric Mapping short course

## Course 4:

## Multiple comparison problem

 \& levels of inferenceChristophe Phillips, Ir PhD
GIGA - CRC In Vivo Imaging \&
GIGA - In Silico Medicine
image data



Statistical Parametric Map


General Linear Model

- model fitting
- statistic image
corrected $p$-values


correction for multiple comparisons




## Content

- Introduction
- Family-wise error rate (FWER)
- False discovery rate (FDR)
- Levels of inference in SPM
- Non-parametric permutation test
- Conclusion


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## Single voxel inference



## Null Hypothesis $\mathbf{H}_{\mathbf{0}}$ : zero activation

Decision rule (threshold) $u$ : determines false positive rate $\alpha$
$\Rightarrow$ Choose $u$ to give acceptable $\alpha$ under $\mathrm{H}_{0}$

Null distribution of test statistic T

$$
\alpha=p\left(t>u \mid H_{0}\right)
$$

## Classical hypothesis testing...

- Null hypothesis H
- test statistic
- null distributions
- Hypothesis test
- control Type I error
- incorrectly reject $H$
- test level $\alpha$
- $\operatorname{Pr}($ "reject" $H \mid H) \leq \alpha$
- $p$-value
- min $\alpha$ at which $H$ rejected
- $\operatorname{Pr}(T \geq t \mid H)$
- characterising "surprise"
$t$-distribution, 32 df .

$F$-distribution, $(10,32) \mathrm{df}$.



## Sensitivity \& specificity

ACTION

|  |  | Don't <br> reject | Reject |
| :---: | :---: | :--- | :--- |
|  | $H_{0}$ true | True <br> Negative | False <br> Positive |
|  | $H_{0}$ false | False <br> Negative | True <br> Positive |

Sensitivity $=T P /(T P+F N)=\beta$
Specificity $=T N /(T N+F P)=1-\alpha$
$\mathrm{FP}=$ Type I error or 'error'
FN = Type II error
$\alpha=\mathrm{p}$-value/FP rate/error rate/significance level
$\beta=$ power

## Multiple tests



Signal


## Multiple tests



If we have 100000 voxels, $\alpha=0.05$
$\Rightarrow \mathbf{5 0 0 0}$ false positive voxels.
This is clearly undesirable!
Need to define a null hypothesis for a collection of tests.

Noisy data


Use of 'uncorrected' $p$-value, $\alpha=0.1$

11.3\%

$12.5 \% \quad 10.8 \% \quad 11.5 \% \quad 10.0 \% \quad 10.7 \% \quad 11.2 \% \quad 10.2 \%$ Percentage of Null Pixels that are False Positives

## Assessing statistics images

## Where's the signal?

High Threshold


Good Specificity
Poor Power
(risk of false negatives)

Med. Threshold


Poor Specificity (risk of false positives)

Good Power

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- Introduction
- Family-wise error rate (FWER)
- Family-wise Null hypothesis
- Bonferroni correction
- Random Field Theory
- False discovery rate (FDR)
- Levels of inference in SPM
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## Family-Wise Null Hypothesis

## Family-Wise Null Hypothesis: Activation is zero everywhere

If we reject a voxel null hypothesis at any voxel, we reject the family-wise Null hypothesis

A FP anywhere in the image gives a Family Wise Error (FWE)
Family-Wise Error rate (FWER) = 'corrected' $p$-value


## Bonferroni correction

The Family-Wise Error rate (FWER), $\alpha_{\text {FWE, }}$ for a family of $N$ tests follows the inequality:

$$
\alpha_{F W E} \leq N \alpha
$$

where $\alpha$ is the test-wise error rate.
Therefore, to ensure a particular FWER choose:

$$
\alpha=\frac{\alpha_{F W E}}{N}
$$

This correction does not require the tests to be independent but becomes very stringent if dependence.

## Bonferroni correction, example

- Experiment with $\mathrm{N}=100000$ independent voxels and 40 d.f.
- $\mathrm{v}=$ unknown corrected probability threshold,
- find $v$ such that family-wise error rate $\alpha=0.05$
- Bonferroni correction:
- probability that all tests are below the threshold,
- use $v=\alpha / N$
- here $v=0.05 / 100000=0.0000005$
$\Rightarrow$ threshold $t=5.77$
- Interpretation:

Bonferroni procedure gives a corrected p-value,
i.e. for a $t$ statistics $=5.77$,

- uncorrectd $p$ value $=0.0000005$
- corrected $p$ value $=0.05$


## Bonferroni \& independent observations



100 by 100 voxels.
10000 independent measures
Fix the $P$ FWE $=0.05, z$ threshold ?
Bonferroni:

$$
\begin{aligned}
v=0.05 / 10000 & =0.000005 \\
& \Rightarrow \text { threshold } z
\end{aligned}=4.42
$$



100 by 100 voxels.
100 independent measures
Fix the $P F W E=0.05, z$ threshold ?
Bonferroni:
$v=0.05 / 100=0.0005$
$\Rightarrow$ threshold $z=3.29$
$v=\alpha / n_{i}$ where $n_{i}$ is the number of independent observations.

## Bonferroni \& independent observations



100 by 100 voxels.
10000 independent measures
Fix the PFWE $=0.05, z$ threshold ?


100 by 100 voxels. How many independent measures ???

Bonferroni:

$$
\begin{aligned}
v=0.05 / 10000 & =0.000005 \\
& \Rightarrow \text { threshold } z=4.42
\end{aligned}
$$

## Random Field Theory

$\Rightarrow$ Consider a statistic image as a discretisation of a continuous underlying random field.
$\Rightarrow$ Use results from continuous random field theory.


## RFT and Euler Characteristic

Euler Characteristic $\chi_{u}$ :

- Topological measure $\chi_{u}=$ \# blobs - \# holes
- at high threshold $u$ :
$\chi_{u}=$ \# blobs


$$
\begin{aligned}
F W E R= & p(F W E) \\
& \approx E\left[\chi_{u}\right]
\end{aligned}
$$

## Euler characteristic...

Threshold z-map


Smoothed image thresholded at $Z>2.5$


Threshold z-map at 2.75



## Expected Euler Characteristic

2D Gaussian Random Field

$100 \times 100$ Gaussian Random Field with FWHM $=10$ smoothing $\alpha_{F W E}=0.05 \Rightarrow u_{R F T}=3.8$
( $u_{\text {BONF }}=4.42, u_{\text {uncorr }}=1.64$ )


## Smoothness

## Smoothness parameterised in terms of FWHM:

Size of Gaussian kernel required to smooth i.i.d. noise to have same smoothness as observed null (standardized) data.

## RESELS (Resolution Elements):

1 RESEL $=F W H M_{x} F W H M_{y} F W H M_{z}$
RESEL Count $R=$ volume of search region in units of smoothness


Eg: 10 voxels, 2.5 FWHM, 4 RESELS

The number of resels is similar, but not identical to the number independent observations.
Smoothness estimated from spatial derivatives of standardised residuals:
Yields an RPV image containing local roughness estimation.


## RFT intuition

Corrected $p$-value for statistic value $t$

$$
\begin{aligned}
p_{c} & =p(\max T>t) \\
& \approx E\left[\chi_{t}\right] \\
& \propto \lambda(\Omega)|\Lambda|^{1 / 2} t \exp \left(-t^{2} / 2\right)
\end{aligned}
$$

- Statistic value $t$ increases ?
- $p_{c}$ decreases (better signal)
- Search volume increases $(\lambda(\Omega) \uparrow)$ ?
- $p_{c}$ increases (more severe correction)
- Smoothness increases $\left(|\Lambda|^{1 / 2} \downarrow\right)$ ?
- $p_{c}$ decreases (less severe correction)


## RFT, unified theory

## General form for expected Euler characteristic

$t, F \& \chi^{2}$ fields $\cdot$ restricted search regions $~ D$ dimensions

$$
E\left[\chi_{u}(\Omega)\right]=\sum_{d=0}^{D} R_{d}(\Omega) \rho_{d}(u)
$$

$\mathrm{R}_{d}(\Omega)$ : d-dimensional Lipschitz-Killing curvatures of $\Omega$ ( $\approx$ intrinsic volumes):

- function of dimension, space $\Omega$ and smoothness:
$\mathrm{R}_{0}(\Omega)=\chi(\Omega)$ Euler characteristic of $\Omega$
$\mathrm{R}_{1}(\Omega)=$ resel diameter
$R_{2}(\Omega)=$ resel surface area
$\mathrm{R}_{3}(\Omega)=$ resel volume
$\rho_{d}(\mathrm{u}): d$-dimensional EC density of the field - function of dimension and threshold, specific for RF type:
E.g. Gaussian RF:

$$
\begin{aligned}
& \rho_{0}(u)=1-\Phi(u) \\
& \rho_{1}(u)=(4 \ln 2)^{1 / 2} \exp \left(-u^{2} / 2\right) /(2 \pi) \\
& \rho_{2}(u)=(4 \ln 2) \quad \mathrm{u} \quad \exp \left(-u^{2} / 2\right) /(2 \pi)^{3 / 2} \\
& \rho_{3}(u)=(4 \ln 2)^{3 / 2}\left(u^{2}-1\right) \quad \exp \left(-u^{2} / 2\right) /(2 \pi)^{2} \\
& \rho_{4}(u)=(4 \ln 2)^{2} \quad\left(u^{3}-3 u\right) \exp \left(-u^{2} / 2\right) /(2 \pi)^{5 / 2}
\end{aligned}
$$

## Estimated component fields


estimate


Each row is an estimated component field

## Smoothness, PRF, ResEls...

- Smoothness $\sqrt{ }|\Lambda|$
- variance-covariance matrix of partial derivatives (possibly location dependent)

$$
\Lambda=\left(\begin{array}{ccc}
\operatorname{var}\left[\frac{\partial e}{\partial x}\right] & \operatorname{cov}\left[\frac{\partial e}{\partial x}, \frac{\partial e}{\partial y}\right] & \operatorname{cov}\left[\frac{\partial e}{\partial x}, \frac{\partial e}{\partial z}\right] \\
\operatorname{cov}\left[\frac{\partial e}{\partial x}, \frac{\partial e}{\partial y}\right] & \operatorname{var}\left[\frac{\partial e}{\partial y}\right] & \operatorname{cov}\left[\frac{\partial e}{\partial y}, \frac{\partial e}{\partial z}\right] \\
\operatorname{cov}\left[\frac{\partial e}{\partial x}, \frac{\partial e}{\partial z}\right] & \operatorname{cov}\left[\frac{\partial e}{\partial y}, \frac{\partial e}{\partial z}\right] & \operatorname{var}\left[\frac{\partial e}{\partial z}\right]
\end{array}\right)
$$

- Point Response Function PRF

- Full Width at Half Maximum FWHM Approximate the peak of the Covariance function with a Gaussian
- Gaussian PRF
- $\Sigma$ - kernel var/cov matrix
- ACF $2 \Sigma$
- $\Lambda=(2 \Sigma)^{-1}$
$\Rightarrow F W H M \mathrm{f}=\sigma \sqrt{ }(8 \ln (2))$
$-\Sigma=\left[\begin{array}{lll}f_{x} & 0 & 0 \\ 0 & f_{y} & 0 \\ 0 & 0 & f_{z}\end{array}\right] \sin (2)$
ignoring covariances
$\Rightarrow \sqrt{ }|\Lambda|=(4 \ln (2))^{3 / 2} /\left(f_{x} \times f_{y} \times f_{z}\right)$
- Resolution Element (ResEl)
- Resel dimensions ( $f_{x} \times f_{y} \times f_{z}$ )
$-R_{3}(\Omega)=\lambda(\Omega) /\left(f_{x} \times f_{y} \times f_{z}\right)$
if strictly stationary

$$
\begin{aligned}
& \mathrm{E}\left[\chi\left(\mathrm{~A}_{u}\right)\right]=\mathrm{R}_{3}(\Omega)(4 \ln (2))^{3 / 2}\left(u^{2}-1\right) \exp \left(-u^{2} / 2\right) \\
& \approx \mathrm{R}_{3}(\Omega)(1-\Phi(u)) \\
& \text { for high thresholds } u
\end{aligned}
$$

## RFT assumptions

- The statistic image is assumed to be a good lattice representation of an underlying random field with a multivariate Gaussian distribution.
- These fields are continuous, with an autocorrelation function twice differentiable at the origin.
$>$ The threshold chosen to define clusters is high enough such that the expected EC is a good approximation to the number of clusters.
> The lattice approximation is reasonable, which implies the smoothness is relatively large compared to the voxel size.
$>$ The errors of the specified statistical model are normally distributed, which implies the model is not misspecified.
- Smoothness of the data is unknown and estimated: very precise estimate by pooling over voxels $\Rightarrow$ stationarity assumption.


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## FDR illustration:



Signal


Signal+Noise


Control of Per Comparison Rate at 10\%


Control of Familywise Error Rate at 10\%


FWE
Occurrence of Familywise Error
Control of False Discovery Rate at 10\%


## Benjamini \& Hochberg Procedure

- Select desired limit $\alpha$ on E(FDR)
- Order p -values, $p_{(1)} \leq p_{(2)} \leq \ldots \leq p_{(v)}$
- Let $r$ be largest $i$ such that

$$
p_{(i)} \leq i / V^{*} \alpha
$$

- Reject all hypotheses corresponding to
$p_{(1)}, \ldots, p_{(r)}$.



## B\&H: Varying Signal Extent

$$
p=
$$

$$
z=
$$




Signal Intensity 3.0
Signal Extent 1.0
Noise Smoothness 3.0

## B\&H: Varying Signal Extent

$$
p=
$$

$$
z=
$$




Signal Intensity 3.0 Signal Extent 2.0 Noise Smoothness 3.0

## B\&H: Varying Signal Extent

$$
p=
$$

$$
z=
$$




Signal Intensity 3.0
Signal Extent 3.0
Noise Smoothness 3.0

## B\&H: Varying Signal Extent

$$
p=0.000252 \quad z=3.48
$$




Signal Intensity 3.0 Signal Extent 5.0 Noise Smoothness 3.0

## B\&H: Varying Signal Extent

$$
p=0.001628 \quad z=2.94
$$




Signal Intensity 3.0
Signal Extent 9.5 Noise Smoothness 3.0

## B\&H: Varying Signal Extent

$$
p=0.007157 \quad z=2.45
$$




Signal Intensity 3.0 Signal Extent16.5 Noise Smoothness 3.0

## B\&H: Varying Signal Extent

$$
p=0.019274 \quad z=2.07
$$




Signal Intensity 3.0 Signal Extent25.0 Noise Smoothness 3.0

## Benjamini \& Hochberg: Properties

- Adaptive
- Larger the signal, the lower the threshold
- Larger the signal, the more false positives
- False positives constant as fraction of rejected tests
- Not a problem with imaging' s sparse signals
- Smoothness OK
- Smoothing introduces positive correlations


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## Topological inference



Peak level inference

人ұ!suəұu!

## Topological feature:

 Peak height
## Topological inference

## Cluster level inference

त
$\stackrel{\text { n }}{ \pm}$
$\pm$
$\pm$

## Topological feature:

 Cluster extent$u_{\text {clus }}:$ cluster-forming threshold



You MUST use a sufficiently high clusterforming threshold $\mathrm{u}_{\text {clus }}$ i.e. $\mathrm{p}_{\text {unc }}<.001$

## Topological inference

## Set level inference



## Peak, cluster \& set level inference

## Sensitivity



[^0]Regional specificity

Peak level test: height of local maxima

Cluster level test: spatial extent above u

Set level test: number of clusters above u


## Levels of inference...

Voxel-level
$\mathrm{P}(\mathrm{c} \geq 1 \mid \mathrm{n} \geq 0, \mathrm{t} \geq 4.37)=0.048$ (corrected) $\mathrm{P}(t \geq 4.37)=1-\Phi\{4.37\}<0.001$ (uncorrected)

$\mathrm{P}(\mathrm{c} \geq 1 \mid \mathrm{n} \geq 82, \mathrm{t} \geq 3.09)=0.029$ (corrected) $\mathrm{P}(\mathrm{n} \geq 82 \mid \mathrm{t} \geq 3.09)=0.019$ (uncorrected)

## Omnibus

$\mathrm{P}(\mathrm{c} \geq 7 \mid \mathrm{n} \geq 0, \mathrm{t} \geq 3.09)=0.031$

Set-level
$\mathrm{P}(\mathrm{c} \geq 3 \mid \mathrm{n} \geq 12, \mathrm{t} \geq 3.09)=0.019$

Parameters

| u | -3.09 |
| :--- | :--- |
| k | -12 voxels |
| S | $-32^{3}$ voxels |
| $F W H M$ | -4.7 voxels |
| D | -3 |

## Small volume correction

If one has some a priori idea of where an activation should be, one can pre-specify a small search space and make the appropriate correction instead of having to control for the entire search space

- mask defined by (probabilistic) anatomical atlases
- mask defined by separate "functional localisers"
- mask defined by orthogonal contrasts
- search volume around previously reported coordinates



## Small Volume Correction

SVC = correction for multiple comparison in a user's defined volume 'of interest'.

Shape and size of volume become important for small or oddly shaped volume!

Example of SVC (900 voxels)

- compact volume: samples from maximum 16 resels
- spread volume: sample from up to 36 resels
$\Rightarrow$ threshold higher for spread volume than compact volume.



## Small volume correction, topology

Table 3. Representative examples of resel counts and critical values.

|  | Vol. | Resel counts |  |  |  |  | $t$ for $\mathrm{P}(M \geq t)=$ |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | :---: |
| Search region $V$ | $(\mathrm{cc})$ | $R_{0}(V)$ | $R_{1}(V)$ | $R_{2}(V)$ | $R_{3}(V)$ | 0.10 | 0.05 | 0.01 |  |
| Single voxel | 0 | 1 | 0 | 0 | 0 | 1.28 | 1.64 | 2.33 |  |
| Head Of Caudate | 7 | 0 | 6.18 | 4.63 | 0.65 | 2.75 | 3.02 | 3.55 |  |
| Putamen | 12 | 1 | 7.32 | 6.80 | 1.18 | 2.89 | 3.15 | 3.66 |  |
| Globus Pallidus | 3 | 0 | 4.03 | 2.29 | 0.24 | 2.49 | 2.78 | 3.35 |  |
| Thalamus | 11 | 1 | 4.94 | 5.14 | 1.13 | 2.79 | 3.05 | 3.59 |  |
| Anterior Cingulate Gyrus | 9 | 1 | 8.20 | 5.79 | 0.86 | 2.86 | 3.11 | 3.63 |  |
| Posterior Cingulate Gyrus | 6 | 1 | 5.32 | 3.85 | 0.58 | 2.70 | 2.97 | 3.51 |  |
| Cingulate Gyri | 15 | 0 | 12.89 | 9.63 | 1.44 | 3.03 | 3.27 | 3.77 |  |
| Superior Frontal Gyrus | 80 | 1 | 15.64 | 25.69 | 8.97 | 3.38 | 3.60 | 4.07 |  |
| Middle Frontal Gyrus | 57 | 1 | 14.89 | 21.14 | 6.23 | 3.31 | 3.53 | 4.00 |  |
| Inferior Frontal Gyrus | 37 | 1 | 11.22 | 14.25 | 4.06 | 3.17 | 3.41 | 3.89 |  |
| Precentral Gyrus | 32 | 1 | 12.30 | 14.23 | 3.40 | 3.16 | 3.40 | 3.88 |  |
| Frontal Gyri | 207 | 1 | 19.30 | 53.39 | 23.63 | 3.63 | 3.84 | 4.28 |  |
| Occipital Lobe | 65 | -1 | 10.68 | 23.11 | 7.17 | 3.32 | 3.55 | 4.02 |  |
| 4mm shell | 254 | 2 | 0.54 | 207.27 | 15.88 | 3.85 | 4.04 | 4.45 |  |
| Whole brain | 1294 | 1 | 20.43 | 107.09 | 153.42 | 4.05 | 4.23 | 4.63 |  |

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## Non-parametric permutation test

- Parametric methods
- Assume distribution of statistic under null hypothesis

- Nonparametric methods
- Use data to find distribution of statistic under null hypothesis
- Any statistic!


Nonparametric Null Distribution

## Permutation Test : Toy Example

- Data from V1 voxel in visual stim. experiment A: Active, flashing checkerboard B: Baseline, fixation 6 blocks, ABABAB Just consider block averages...

| A | B | A | B | A | B |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 103.00 | 90.48 | 99.93 | 87.83 | 99.76 | 96.06 |

- Null hypothesis $H_{0}$
- No experimental effect, A \& B labels arbitrary
- Statistic
- Mean difference


## Permutation Test : Toy Example

- Under $H_{0}$
- Consider all equivalent relabelings

| AAABBB | ABABAB | BAAABB | BABBAA |
| :--- | :--- | :--- | :--- |
| AABABB | ABABBA | BAABAB | BBAAAB |
| AABBAB | ABBAAB | BAABBA | BBAABA |
| AABBBA | ABBABA | BABAAB | BBABAA |
| ABAABB | ABBBAA | BABABA | BBBAAA |

## Permutation Test : Toy Example

- Under $H_{0}$
- Consider all equivalent relabelings
- Compute all possible statistic values

| AAABBB | 4.82 | ABABAB | 9.45 | BAAABB | $\mathbf{- 1 . 4 8}$ | BABBAA | -6.86 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| AABABB | -3.25 | ABABBA | 6.97 | BAABAB | 1.10 | BBAAAB | 3.15 |
| AABBAB | -0.67 | ABBAAB | 1.38 | BAABBA | $\mathbf{- 1 . 3 8}$ | BBAABA | 0.67 |
| AABBBA | $-\mathbf{- 3 . 1 5}$ | ABBABA | $\mathbf{- 1 . 1 0}$ | BABAAB | -6.97 | BBABAA | 3.25 |
| ABAABB | 6.86 | ABBBAA | 1.48 | BABABA | $\mathbf{- 9 . 4 5}$ | BBBAAA | $\mathbf{- 4 . 8 2}$ |

## Permutation Test : Toy Example

- Under $H_{0}$
- Consider all equivalent relabelings
- Compute all possible statistic values
- Find 95\%ile of permutation distribution

| AAABBB | 4.82 | ABABAB | 9.45 | BAAABB | $\mathbf{- 1 . 4 8}$ | BABBAA | -6.86 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| AABABB | -3.25 | ABABBA | 6.97 | BAABAB | 1.10 | BBAAAB | 3.15 |
| AABBAB | -0.67 | ABBAAB | 1.38 | BAABBA | $\mathbf{- 1 . 3 8}$ | BBAABA | 0.67 |
| AABBBA | -3.15 | ABBABA | $\mathbf{- 1 . 1 0}$ | BABAAB | -6.97 | BBABAA | 3.25 |
| ABAABB | 6.86 | ABBBAA | 1.48 | BABABA | $\mathbf{- 9 . 4 5}$ | BBBAAA | $\mathbf{- 4 . 8 2}$ |

## Permutation Test : Toy Example

- Under $H_{0}$
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## Permutation Test : Toy Example

- Under $H_{0}$
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- Find 95\%ile of permutation distribution

| AAABBB | 4.82 | ABABAB | 9.45 | BAAABB | $\mathbf{- 1 . 4 8}$ | BABBAA | -6.86 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| AABABB | -3.25 | ABABBA | 6.97 | BAABAB | 1.10 | BBAAAB | 3.15 |
| AABBAB | -0.67 | ABBAAB | 1.38 | BAABBA | $\mathbf{- 1 . 3 8}$ | BBAABA | 0.67 |
| AABBBA | -3.15 | ABBABA | -1.10 | BABAAB | -6.97 | BBABAA | 3.25 |
| ABAABB | 6.86 | ABBBAA | 1.48 | BABABA | $\mathbf{- 9 . 4 5}$ | BBBAAA | $\mathbf{- 4 . 8 2}$ |

## Controlling FWER: Permutation Test

- Parametric methods
- Assume distribution of max statistic under null hypothesis

- Nonparametric methods
- Use data to find distribution of max statistic under null hypothesis
- Again, any max statistic!


Nonparametric Null Max Distribution

## Permutation Test \& Exchangeability

- Exchangeability is fundamental
- Def: Distribution of the data unperturbed by permutation
- Under $\mathrm{H}_{0}$, exchangeability justifies permuting data
- Allows us to build permutation distribution
- Subjects are exchangeable
- Under $H_{o}$, each subject's A/B labels can be flipped
- Are fMRI scans exchangeable under $\mathrm{H}_{0}$ ?
- If no signal, can we permute over time?


## Permutation Test \& Exchangeability

- fMRI scans are not exchangeable
- Permuting disrupts order, temporal autocorrelation
- Intrasubject fMRI permutation test
- Must decorrelate data, model before permuting
- What is correlation structure?
- Usually must use parametric model of correlation
- E.g. Use wavelets to decorrelate
- Bullmore et al 2001, HBM 12:61-78
- Intersubject fMRI permutation test
- Create difference image for each subject
- For each permutation, flip sign of some subjects


## Permutation Test : Example

- fMRI Study of Working Memory
- 12 subjects, block design Marshuetz et al (2000)
- Item Recognition
- Active:View five letters, 2s pause, view probe letter, respond
- Baseline: View XXXXX, 2s pause,
view Y or N, respond
- Second Level RFX
- Difference image, A-B constructed for each subject
- One sample, smoothed variance $t$ test



## Permutation Test : Example

- Permute!
$-2^{12}=4,096$ ways to flip 12 A/B labels
- For each, note maximum of $t$ image


Permutation Distribution Maximum $t$


Maximum Intensity Projection Thresholded $t$


$u^{\text {Perm }}=7.67$
58 sig. vox.
$t_{11}$ Statistic, Nonparametric Threshold


Test Level vs. $t_{11}$ Threshold

- Compare with Bonferroni
- Compare with parametric RFT
$110,7762 \times 2 \times 2 \mathrm{~mm}$ voxels $5.1 \times 5.8 \times 6.9 \mathrm{~mm}$ FWHM
smoothness 462.9 RESELs

$$
\alpha=0.05 / 110,776
$$

$t_{11}$ Statistic, RF \& Bonf. Threshold

$$
\begin{aligned}
& u^{u^{\mathrm{RF}}=9.87} \\
& u^{\text {Bonf }}=9.80 \\
& 5 \text { sig. vox. }
\end{aligned}
$$

## Generalization: RFT vs Bonf. vs Perm.

|  |  | $t$ Threshold |  |  |
| :--- | :--- | ---: | ---: | ---: |
|  | df | (0.05 Corrected) |  |  |
|  |  | Bonf | Perm |  |
| Verbal Fluency | 4 | 4701.32 | 42.59 | 10.14 |
| Location Switching | 9 | 11.17 | 9.07 | 5.83 |
| Task Switching | 9 | 10.79 | 10.35 | 5.10 |
| Faces: Main Effect | 11 | 10.43 | 9.07 | 7.92 |
| Faces: Interaction | 11 | 10.70 | 9.07 | 8.26 |
| Item Recognition | 11 | 9.87 | 9.80 | 7.67 |
| Visual Motion | 11 | 11.07 | 8.92 | 8.40 |
| Emotional Pictures | 12 | 8.48 | 8.41 | 7.15 |
| Pain: Warning | 22 | 5.93 | 6.05 | 4.99 |
| Pain: Anticipation | 22 | 5.87 | 6.05 | 5.05 |

## RFT vs Bonf. vs Perm.

|  |  | No. Significant Voxels |  |  |
| :--- | :---: | ---: | ---: | ---: | ---: |
| $(0.05$ | Corrected) |  |  |  |

## Content

- Introduction
- Family-wise error rate (FWER)
- False discovery rate (FDR)
- Levels of inference in SPM
- Non-parametric permutation test
- Conclusion


## What we' d like

- Don't threshold, model the signal!
- Signal location?
- Estimates and CI's on ( $x, y, z$ ) location
- Signal magnitude?
- CI's on \% change
- Spatial extent?

- Estimates and CI's on activation volume
- Robust to choice of cluster definition
- ...but this requires an explicit spatial model


## Real-life inference: What we get

- Signal location
- Local maximum - no inference
- Center-of-mass - no inference
- Sensitive to blob-defining-threshold
- Signal magnitude
- Local maximum intensity - P-values (\& CI's)
- Spatial extent
- Cluster volume - P-value, no CI's
- Sensitive to blob-defining-threshold


## FWER vs. FDR

You MUST account for multiplicity
(Otherwise have a fishing expedition)

- FWER
- Very specific, not very sensitive
- FDR
- Less specific, more sensitive
(Sociological calibration still underway)


## Conclusion

- There is a multiple testing problem and corrections must be applied on p-values, possibly for the volume of interest only (see SVC).
- Inference is made about topological features (peak height, spatial extent, number of clusters). Use results from the Random Field Theory. Or permutation tests.
- Control of FWER (probability of a false positive anywhere in the image) for a space of any dimension and shape.


## References

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- And now a little demo!


[^0]:    $\mathrm{L}_{1}>$ spatial extent threshold
    $\mathrm{L}_{2}$ < spatial extent threshold

