

# Introduction à la statistique médicale

## Statistical Parametric Mapping short course

### Course 3:

### General Linear Model, Contrast & Inference

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GIGA – CRC *In Vivo* Imaging &

GIGA – *In Silico* Medicine

# Content

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- **Introduction**
- **Contrast & Inference**
- **Orthogonality issue**
- **Conclusion**

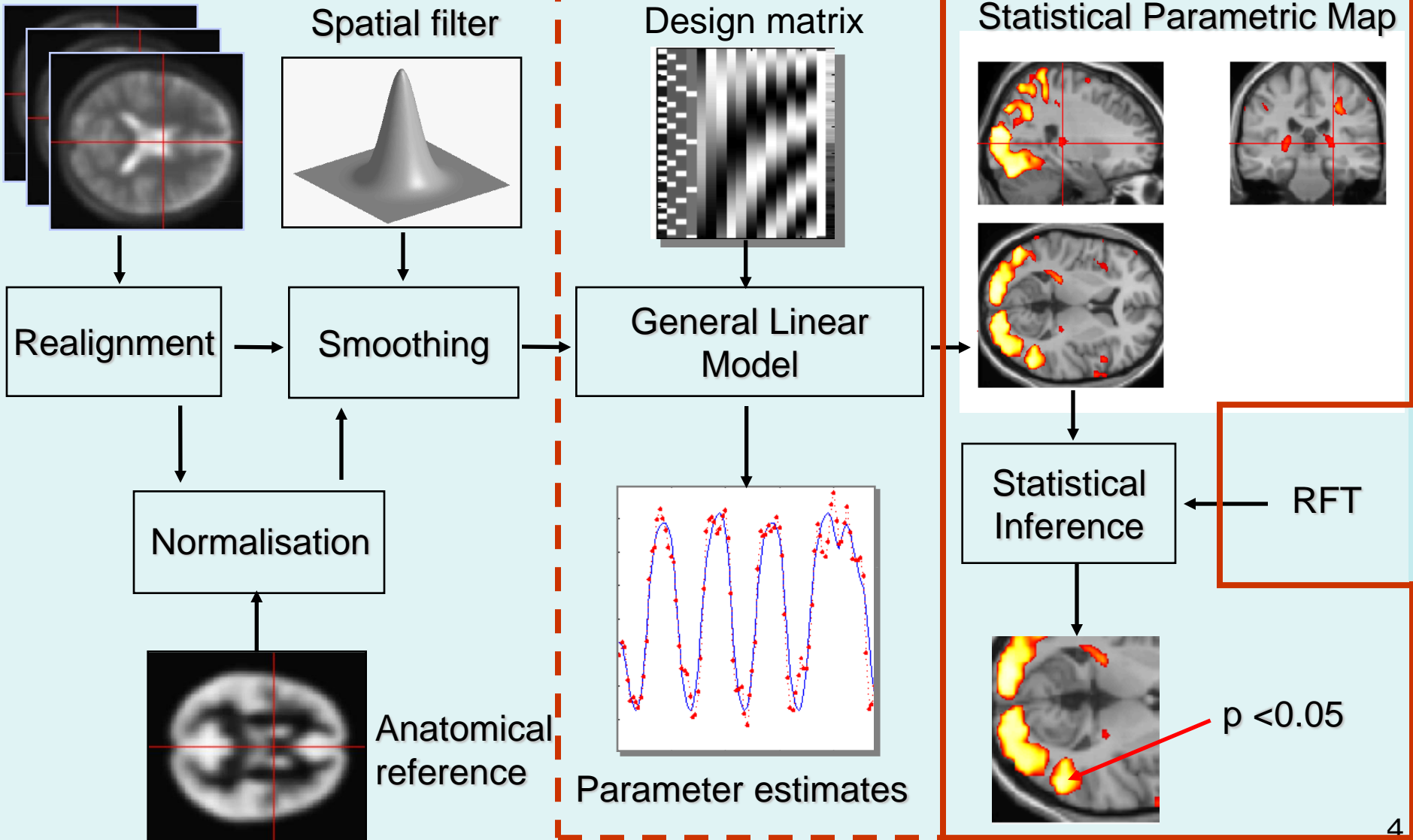
# Content

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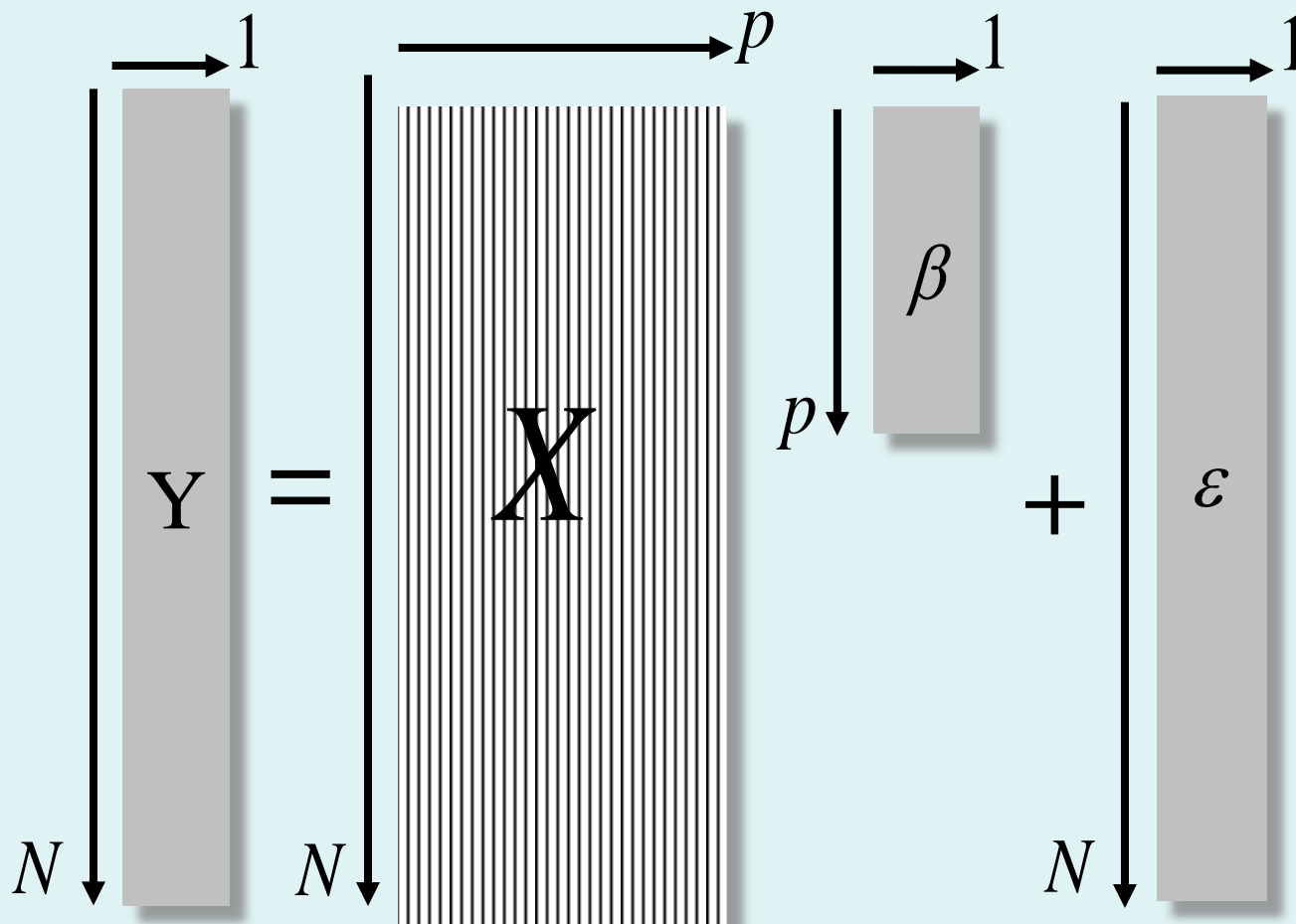
- **Introduction**
  - **Generalized Linear Model**
  - **Estimated parameters**
- **Contrast & Inference**
- **Orthogonality issue**
- **Conclusion**

# SPM work flow

Image time-series



# General Linear Model



$$Y = X\beta + \epsilon$$

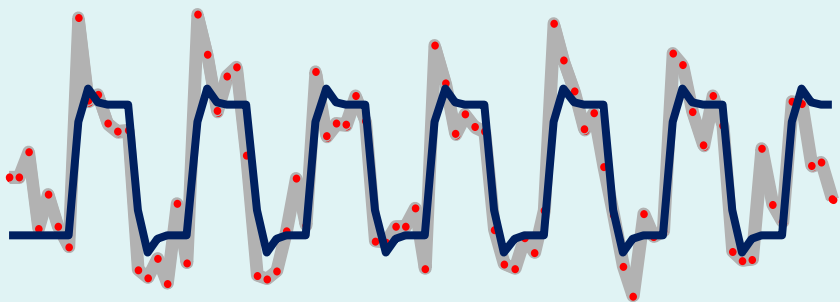
$$\epsilon \sim N(0, \sigma^2 I)$$

This is for a  
SINGLE voxel !  
Design matrix  $X$   
is the same for  
ALL voxels !

$N$ : number of scans  
 $p$ : number of regressors

Model is specified by  
1. Design matrix  $X$   
2. Assumptions about  $\epsilon$

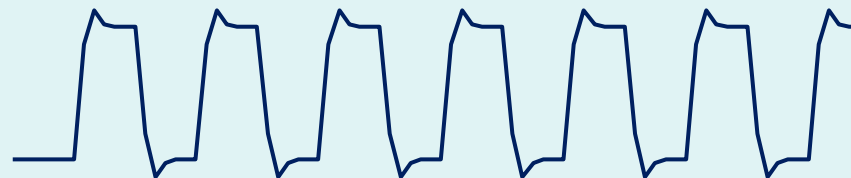
# Estimation of the parameters



**i.i.d. assumptions:**  $\varepsilon \sim N(0, \sigma^2 I)$

**OLS estimates:**  $\hat{\beta} = (X^T X)^{-1} X^T y$

$$\hat{\beta}_1 = 3.9831$$



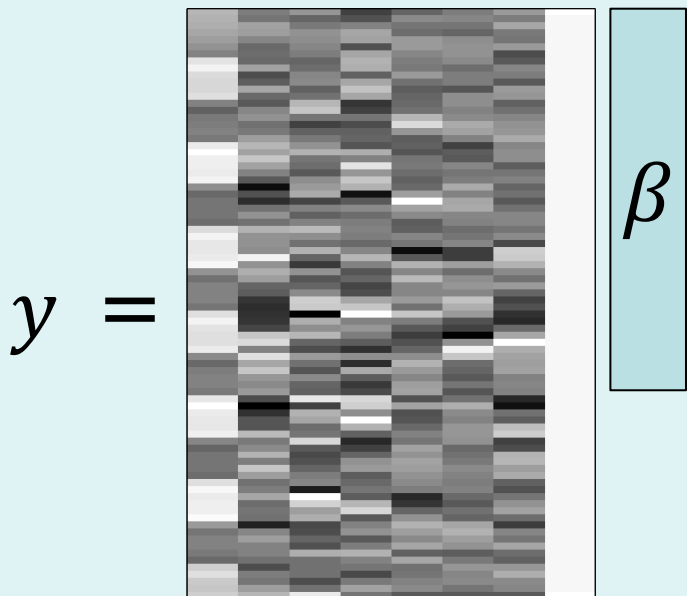
$$\hat{\beta}_{2-7} = \{0.6871, 1.9598, 1.3902, 166.1007, 76.4770, -64.8189\}$$



$$\hat{\beta}_8 = 131.0040$$

+  $\varepsilon$

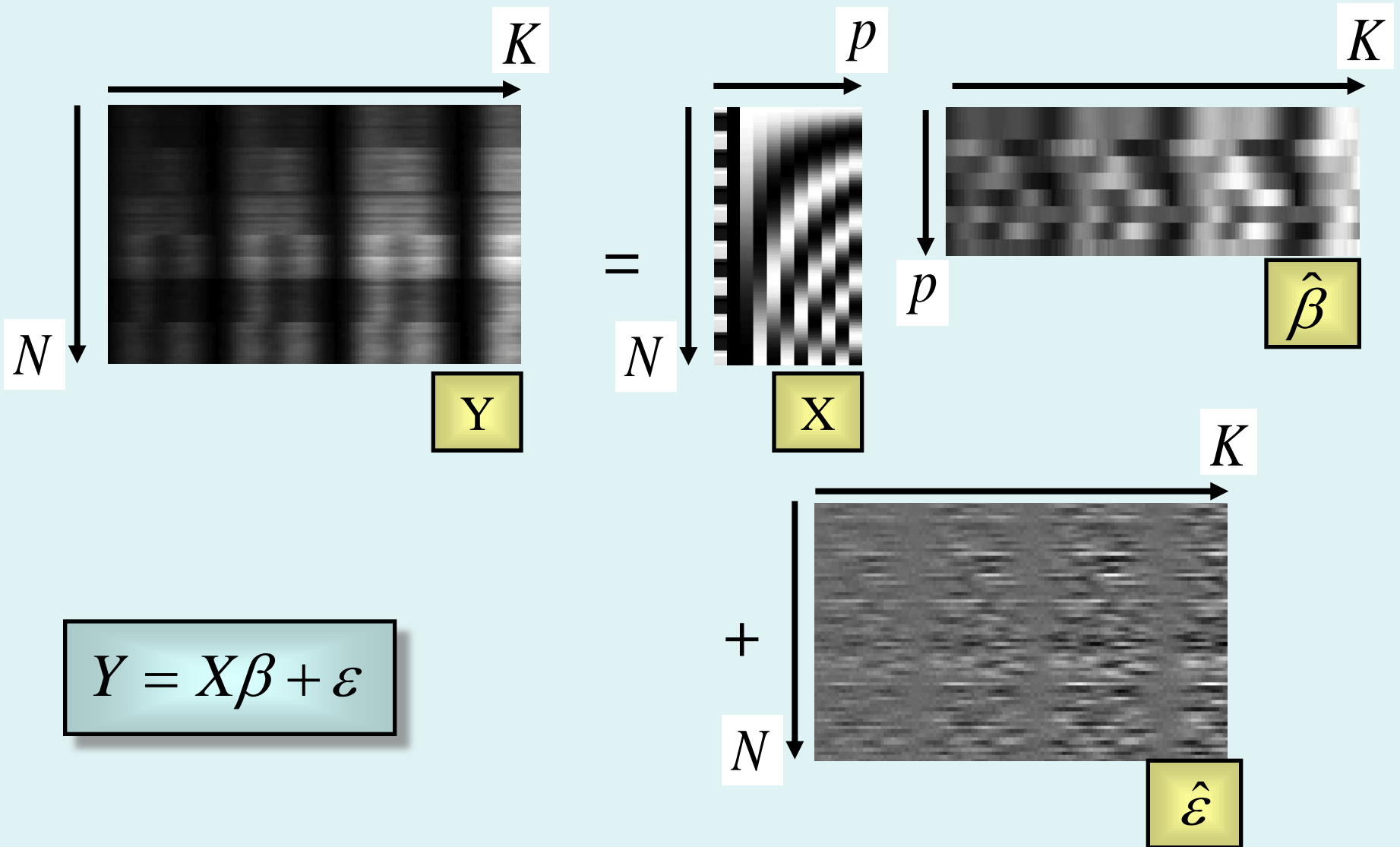
$$\hat{\varepsilon} =$$



$$\hat{\beta} \sim N(\beta, \sigma^2 (X^T X)^{-1})$$

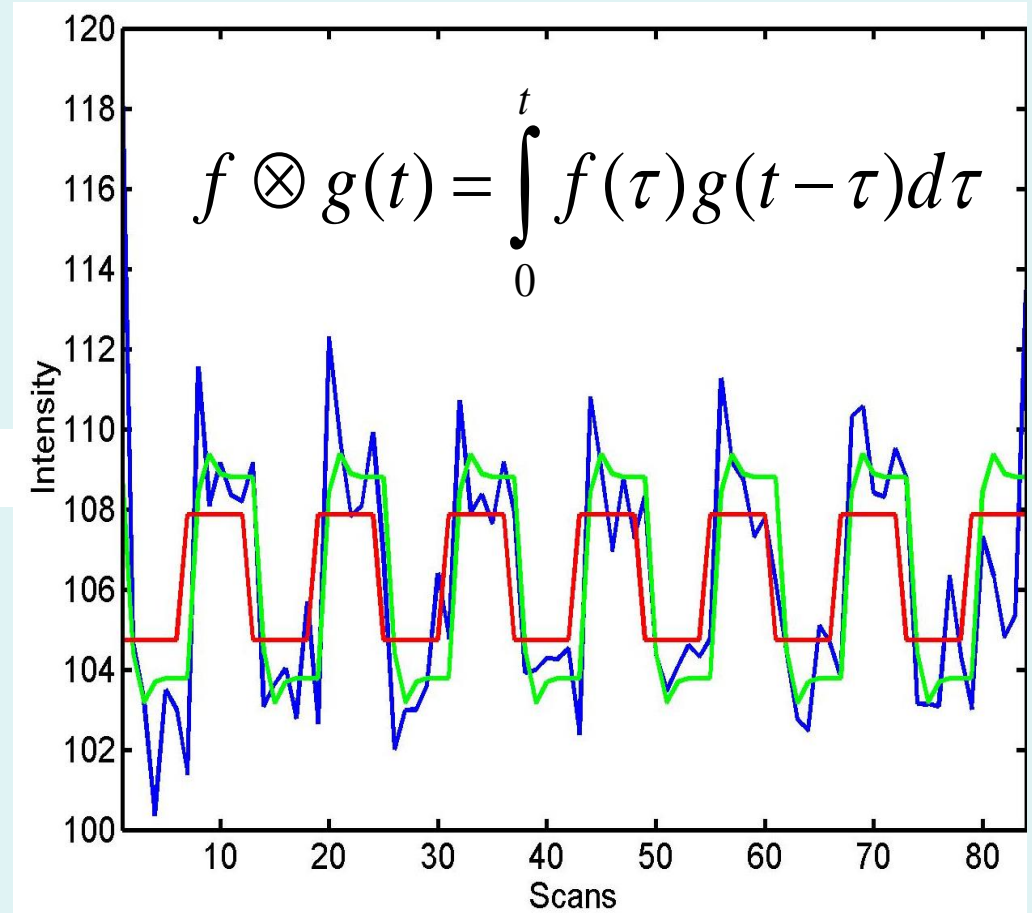
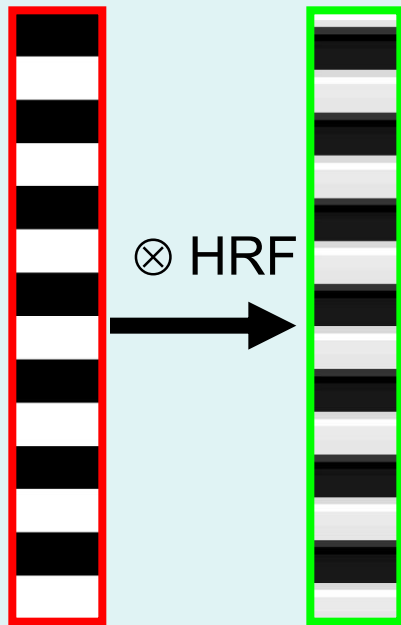
$$\hat{\sigma}^2 = \frac{\hat{\varepsilon}^T \hat{\varepsilon}}{N-p}$$

# GLM & Mass univariate approach



# Convolution model of the BOLD response

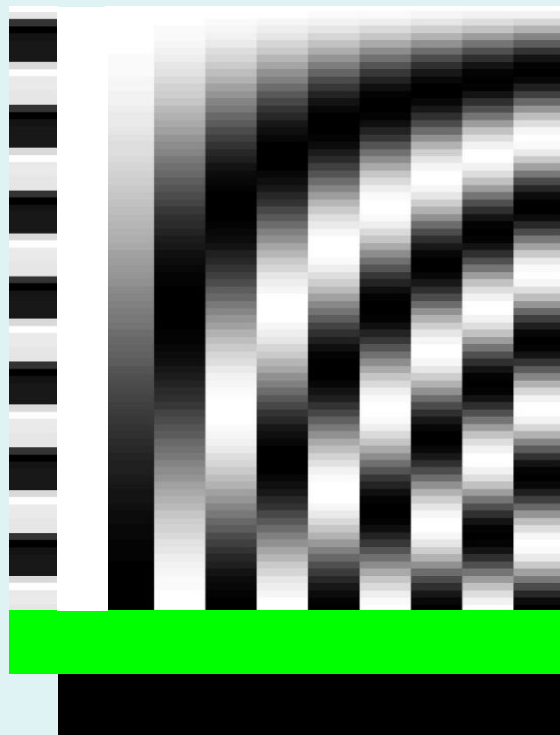
Convolve stimulus function with a canonical hemodynamic response function (HRF):



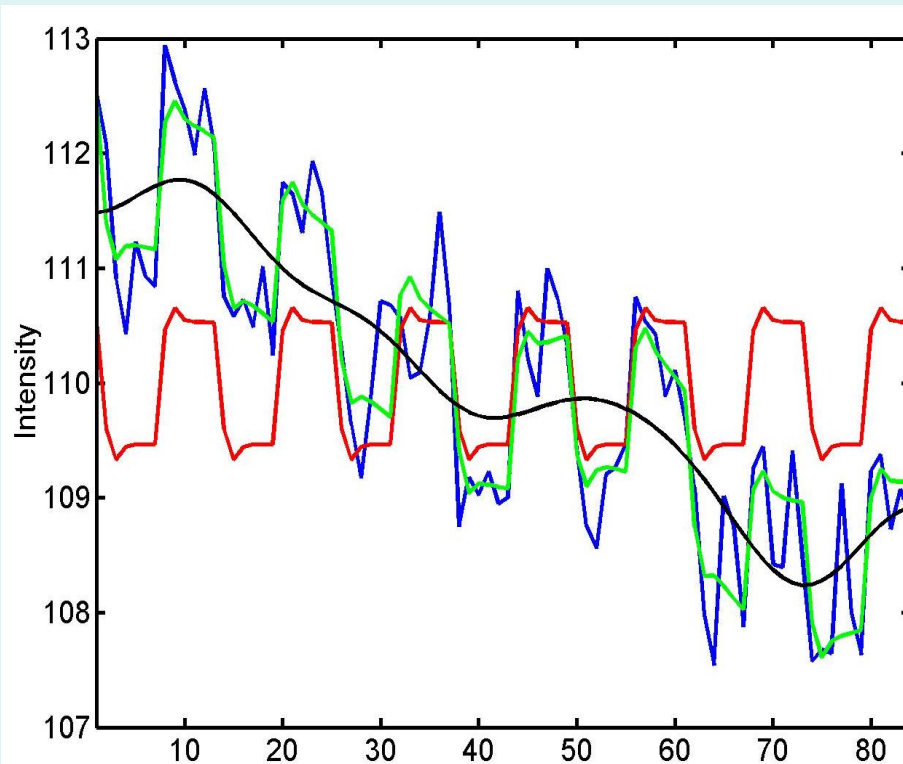


# Low-frequency noise

Solution: High pass filtering



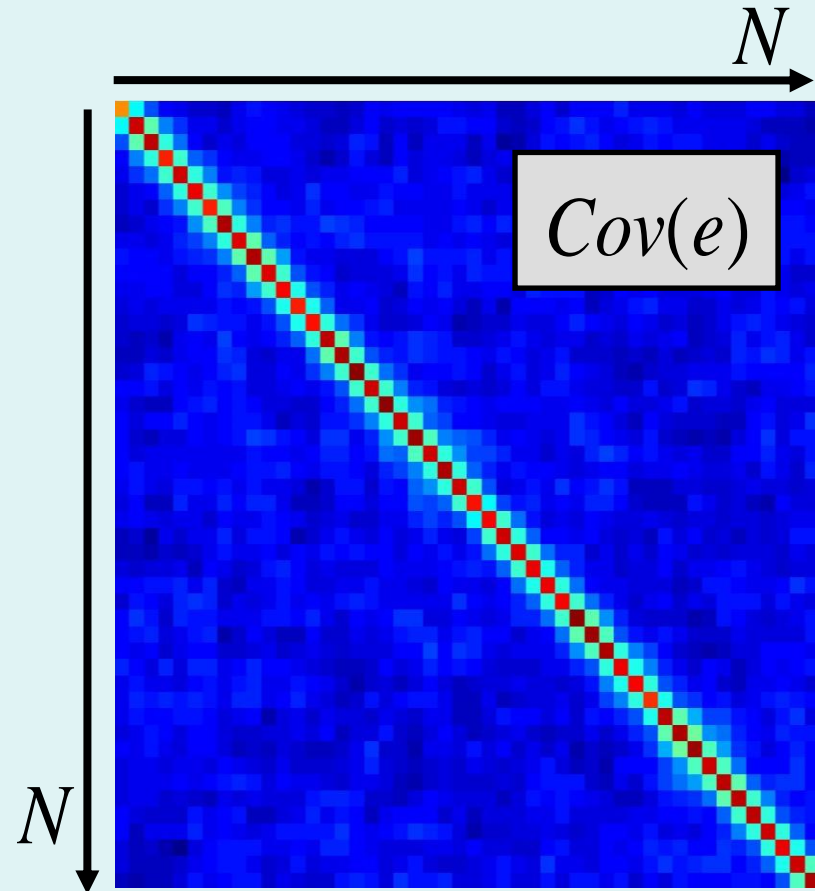
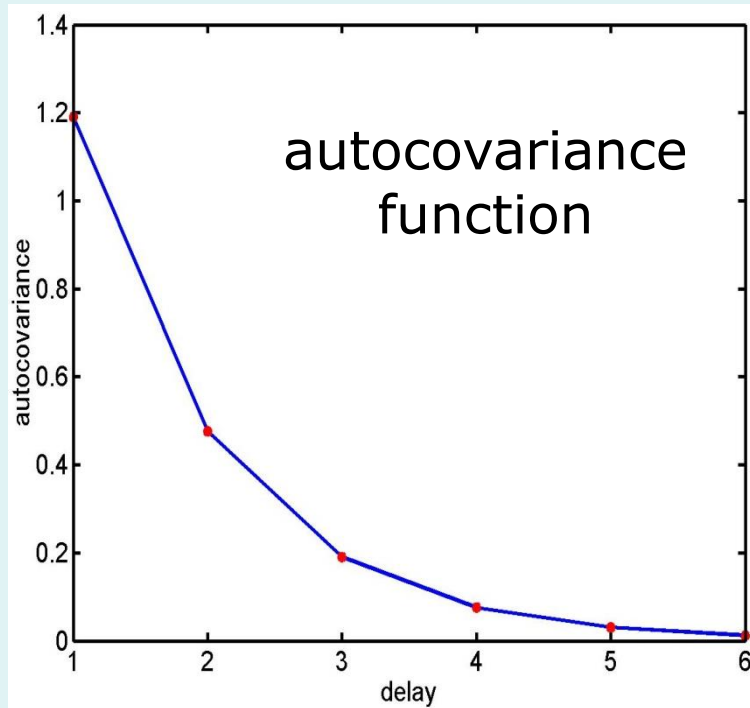
discrete cosine  
transform (DCT)  
set



blue = data  
black = mean + low-frequency drift  
green = predicted response, taking into account low-frequency drift  
red = predicted response, NOT taking into account low-frequency drift

# 3. Serial correlation

~~i.i.d:  $\varepsilon \sim N(0, \sigma^2 I)$~~



# Multiple covariance components

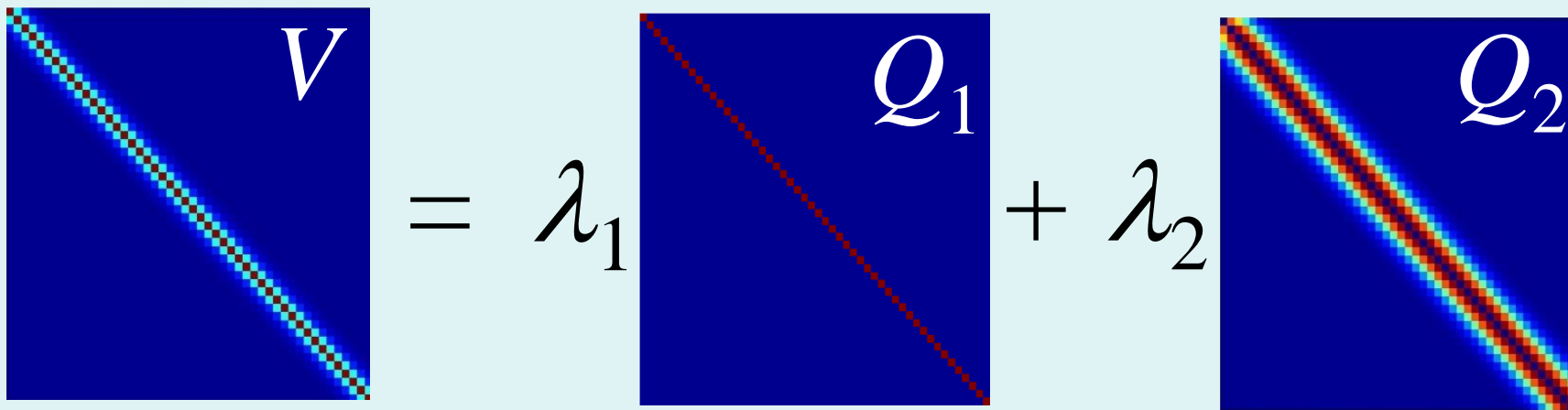
enhanced noise model at voxel  $i$

$$e_i \sim N(0, C_i)$$

$$C_i = \sigma_i^2 V$$

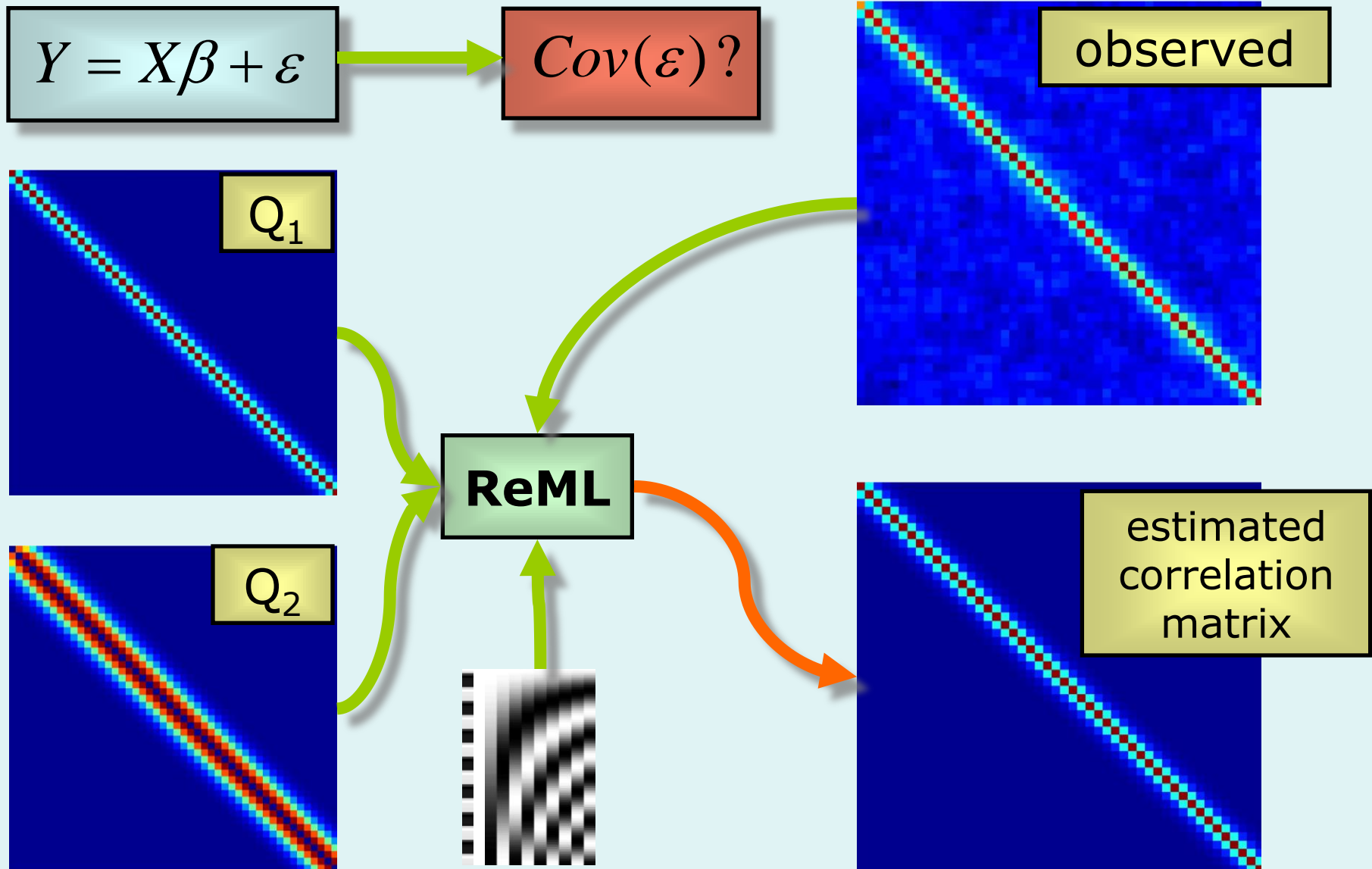
$$V = \sum \lambda_j Q_j$$

error covariance components  
 $Q$  and hyperparameters  $\lambda$

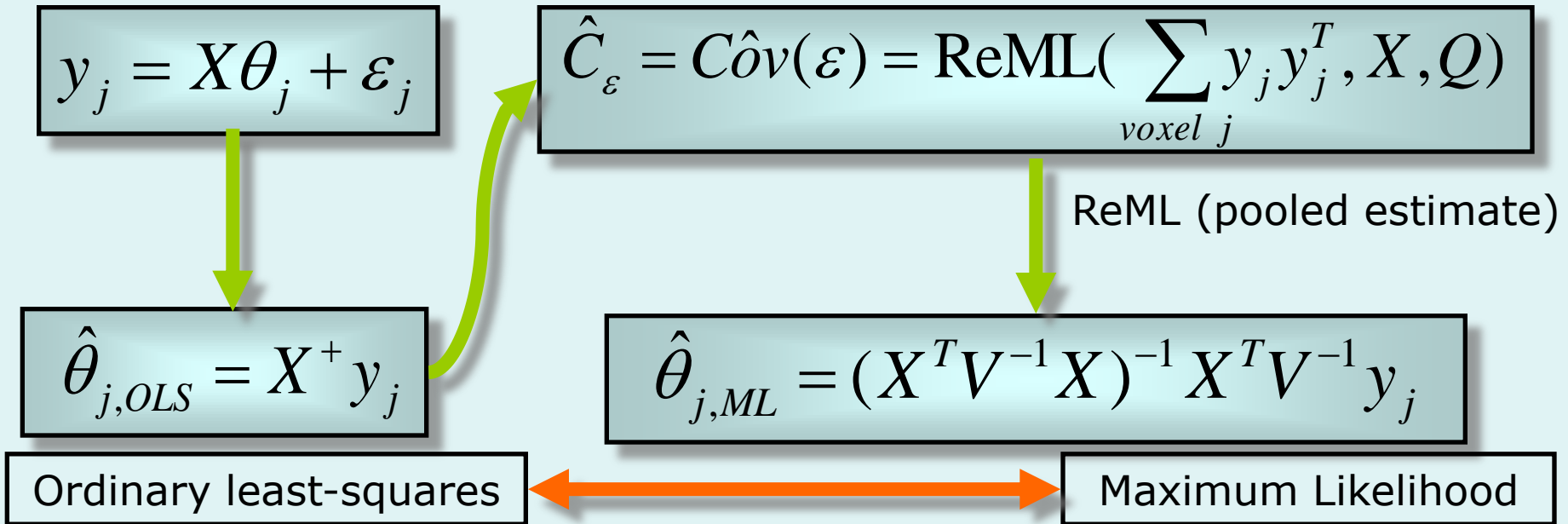


Estimation of hyperparameters  $\lambda$  with ReML (Restricted Maximum Likelihood).

# Restricted Maximul Likelihood



# Estimation in SPM



- 2 passes (first pass for selection of voxels)
- more accurate estimate of  $V$

Assume, at voxel  $j$ :  $C_{\varepsilon,j} = \sigma_j V$

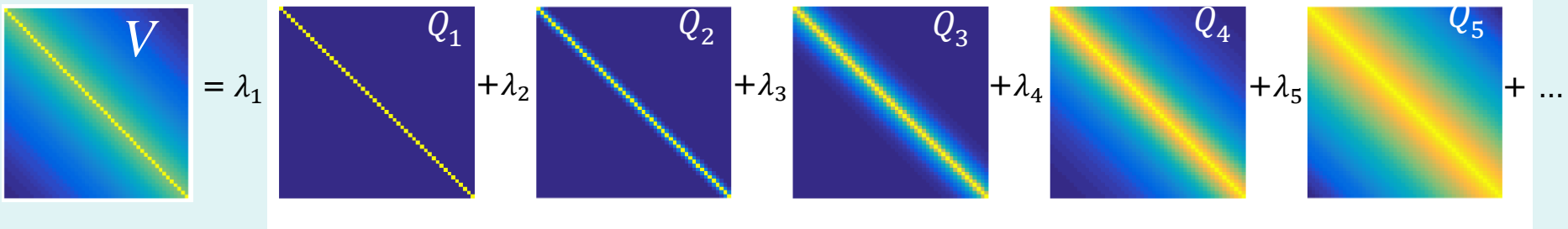
$$t = \frac{c^T \theta}{\text{SE}(c^T \theta)}$$

$$\text{SE}(c^T \theta) = \sqrt{\hat{\sigma}^2 c^T (V^{-1/2} X)^{-1} (V^{-1/2} X)^{-T} c}$$

# Limitations



The AR(1)+white noise model may not be enough for short TR (<1.5 s)



The flexibility of the ReML enables the use of any number of components of any shape

# Content

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- **Introduction**
- **Contrast & Inference**
  - **Hypothesis testing**
  - **Contrast**
  - **t-Test**
  - **F-test**
- **Orthogonality issue**
- **Conclusion**

# Hypothesis testing

To test a hypothesis, we construct “test statistics”.

- **Null Hypothesis  $H_0$**

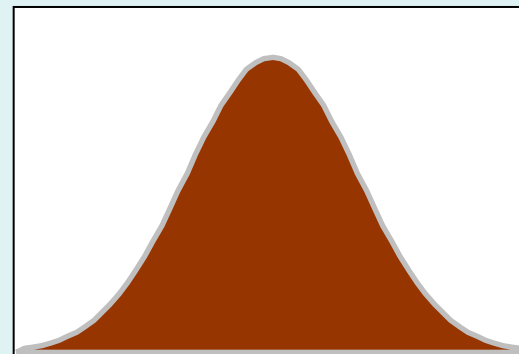
Typically what we want to disprove (no effect).

→ *Alternative Hypothesis  $H_A$*  expresses outcome of interest

- **Test Statistic  $T$**

The test statistic summarises evidence about  $H_0$ .

Typically, test statistic is small in magnitude when the hypothesis  $H_0$  is true and large when false.



Null Distribution of  $T$

→ We need to know the distribution of  $T$  under the null hypothesis.



# Hypothesis testing & inference

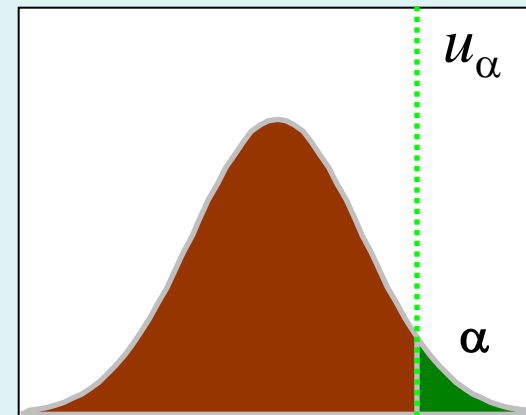
- **Significance level  $\alpha$ :**

Acceptable *false positive rate*  $\alpha$ .

⇒ threshold  $u_\alpha$

Threshold  $u_\alpha$  controls the false positive rate

$$\alpha = p(T > u_\alpha | H_0)$$



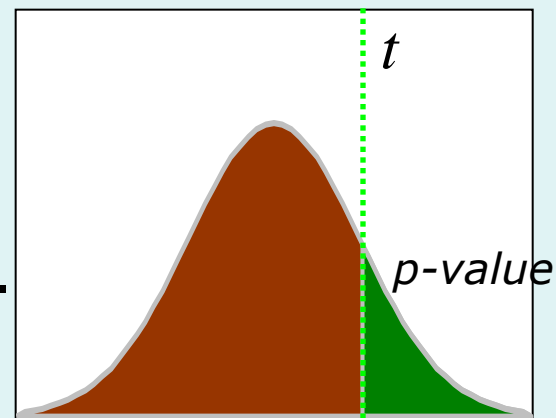
Null Distribution of T

- **Conclusion about the hypothesis:**

We reject the null hypothesis in favour of the alternative hypothesis if  $t > u_\alpha$

- **p-value:**

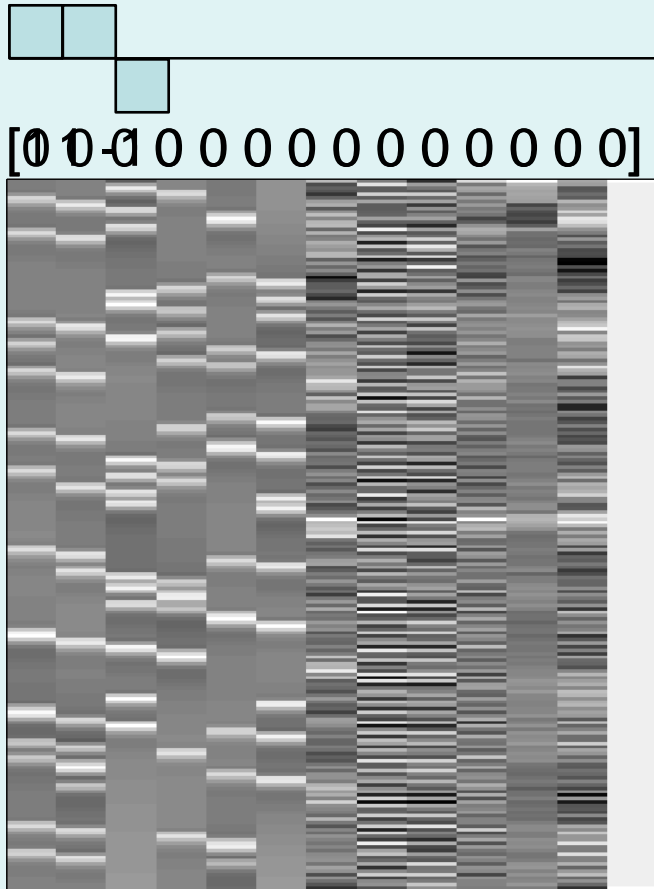
A *p-value* summarises evidence against  $H_0$ . This is the chance of observing value more extreme than  $t$  under the null hypothesis.



Null Distribution of T

$$p(T > t | H_0)$$

# Contrast & effect of interest



A contrast selects a **specific effect of interest**

- a contrast  $c$  is a vector of length  $p$ .
- $c^T \beta$  is a linear combination of regression coefficients  $\beta$ .

$$c = [1 \ 0 \ 0 \ 0 \ \dots]^T$$

$$\begin{aligned} c^T \beta &= \mathbf{1} \times \beta_1 + \mathbf{0} \times \beta_2 + \mathbf{0} \times \beta_3 + \mathbf{0} \times \beta_4 + \dots \\ &= \beta_1 \end{aligned}$$

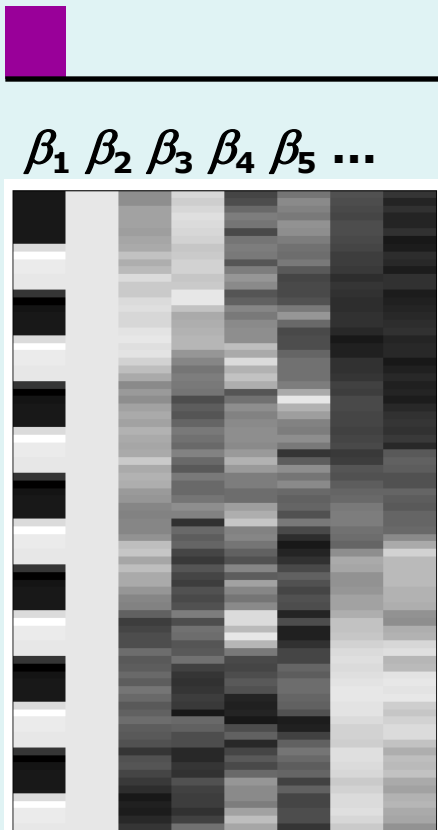
$$c = [0 \ 1 \ -1 \ 0 \ \dots]^T$$

$$\begin{aligned} c^T \beta &= \mathbf{0} \times \beta_1 + \mathbf{1} \times \beta_2 + \mathbf{-1} \times \beta_3 + \mathbf{0} \times \beta_4 + \dots \\ &= \beta_2 - \beta_3 \end{aligned}$$

$$c^T \hat{\beta} \sim N(c^T \beta, \sigma^2 c^T (X^T X)^{-1} c)$$

# t-Test, one dimensional contrast

$$c^T = 1 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0$$



**Question:** box-car amplitude  $> 0$  ?  
=  
 $\beta_1 = c^T \beta > 0$  ?

**Null hypothesis:**

$$H_0: c^T \beta = 0$$

**contrast of  
estimated  
parameters**

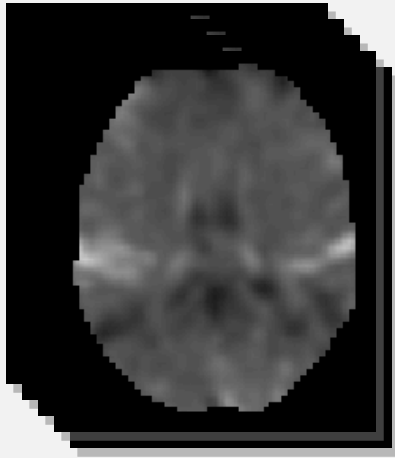
**Test statistic:**

$$T = \frac{\text{contrast of estimated parameters}}{\sqrt{\text{variance estimate}}}$$

$$T = \frac{c^T \hat{\beta}}{\sqrt{\text{var}(c^T \hat{\beta})}} = \frac{c^T \hat{\beta}}{\sqrt{\hat{\sigma}^2 c^T (X^T X)^{-1} c}} \sim t_{N-p}$$

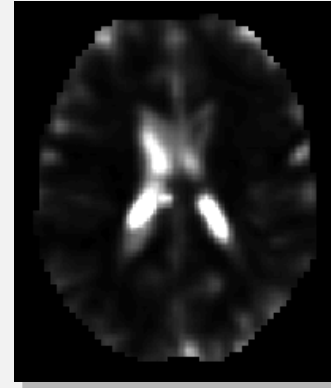
# t-Test in SPM

For a given contrast  $c$ :



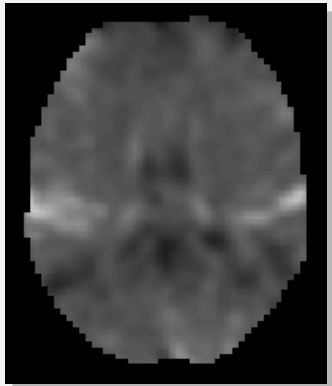
beta\_???? images

$$\hat{\beta} = (X^T X)^{-1} X^T y$$



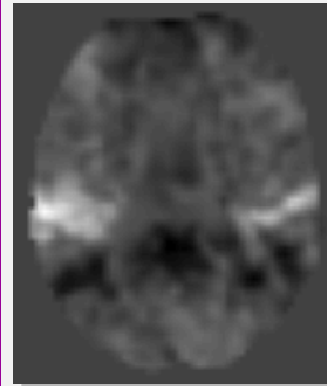
ResMS image

$$\hat{\sigma}^2 = \frac{\hat{\varepsilon}^T \hat{\varepsilon}}{N - p}$$



con\_???? image

$$c^T \hat{\beta}$$

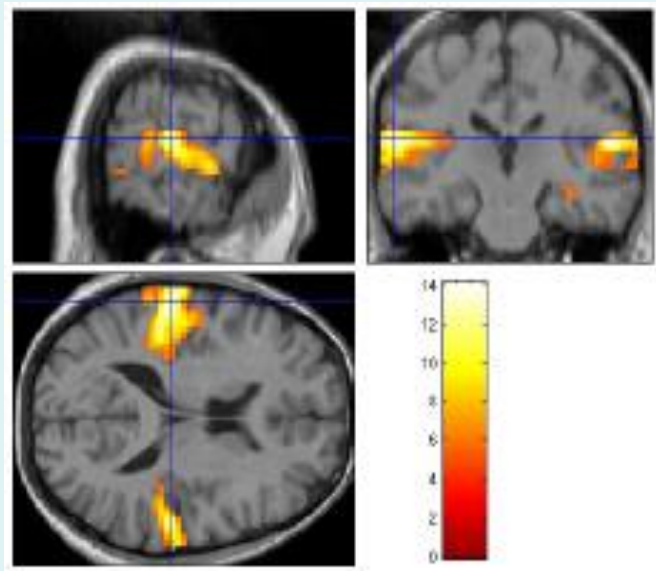


spmT\_???? image

$$\text{SPM}\{t\}$$

# t-Test, simple example

Passive word listening vs. rest



$$c^T = [1 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0]$$

Q: activation during listening ?

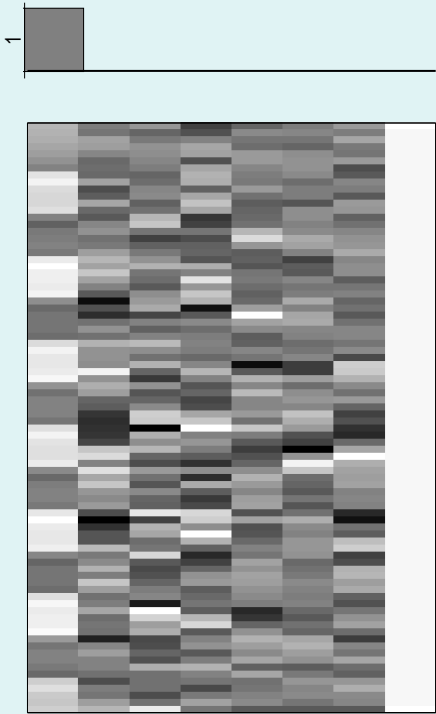
Null hypothesis:  $\beta_1 = 0$

$$t = \frac{c^T \hat{\beta}}{\sqrt{\text{var}(c^T \hat{\beta})}}$$

SPMresults:

Height threshold  $T = 3.2057$   $\{p < 0.001\}$

voxel-level			mm mm mm		
$T$	$(Z)$	$\rho_{\text{uncorrected}}$			
<b>13.94</b>	<b>Inf</b>	<b>0.000</b>	<b>-63</b>	<b>-27</b>	<b>15</b>
12.04	Inf	0.000	-48	-33	12
11.82	Inf	0.000	-66	-21	6
<b>13.72</b>	<b>Inf</b>	<b>0.000</b>	<b>57</b>	<b>-21</b>	<b>12</b>
12.29	Inf	0.000	63	-12	-3
9.89	7.83	0.000	57	-39	6
<b>7.39</b>	<b>6.36</b>	<b>0.000</b>	<b>36</b>	<b>-30</b>	<b>-15</b>
<b>6.84</b>	<b>5.99</b>	<b>0.000</b>	<b>51</b>	<b>0</b>	<b>48</b>
<b>6.36</b>	<b>5.65</b>	<b>0.000</b>	<b>-63</b>	<b>-54</b>	<b>-3</b>
<b>6.19</b>	<b>5.53</b>	<b>0.000</b>	<b>-30</b>	<b>-33</b>	<b>-18</b>
<b>5.96</b>	<b>5.36</b>	<b>0.000</b>	<b>36</b>	<b>-27</b>	<b>9</b>
<b>5.84</b>	<b>5.27</b>	<b>0.000</b>	<b>-45</b>	<b>42</b>	<b>9</b>
<b>5.44</b>	<b>4.97</b>	<b>0.000</b>	<b>48</b>	<b>27</b>	<b>24</b>
<b>5.32</b>	<b>4.87</b>	<b>0.000</b>	<b>36</b>	<b>-27</b>	<b>42</b>



# t-Test, summary

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- $T$ -test is a signal-to-noise measure (ratio of estimate to standard deviation of estimate).

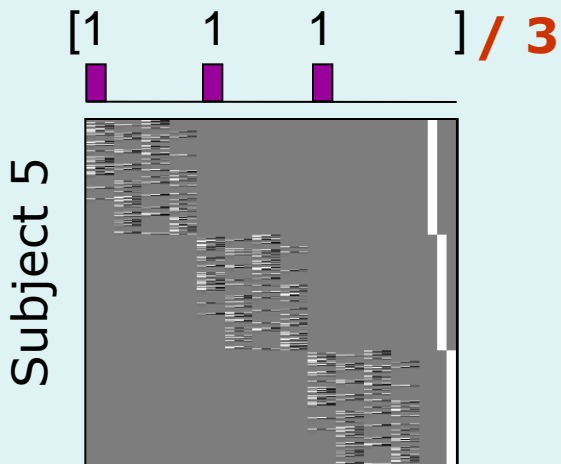
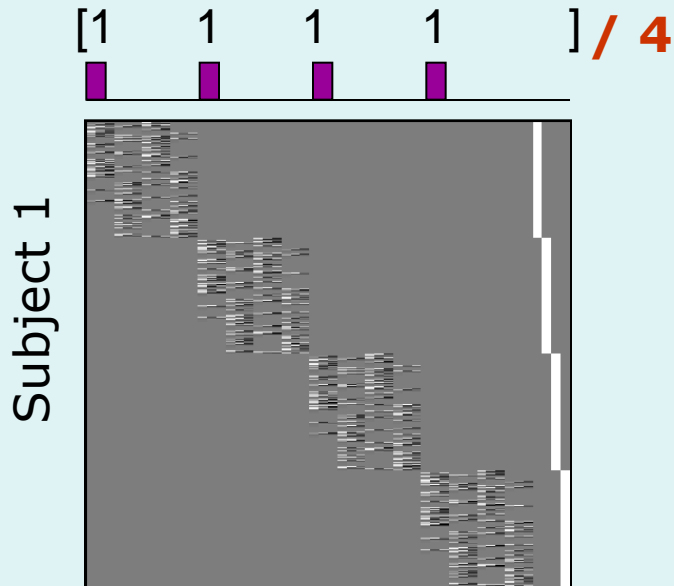
- Alternative hypothesis:

$$H_0: c^T \beta = 0 \quad \text{vs} \quad H_A: c^T \beta > 0$$

- $T$ -contrasts are simple combinations of the betas
- $T$ -statistic *does not depend* on the scaling of the regressors or the scaling of the contrast.

# t-Test, scaling issue

$$T = \frac{c^T \hat{\beta}}{\sqrt{\text{var}(c^T \hat{\beta})}} = \frac{c^T \hat{\beta}}{\sqrt{\hat{\sigma}^2 c^T (X^T X)^{-1} c}}$$



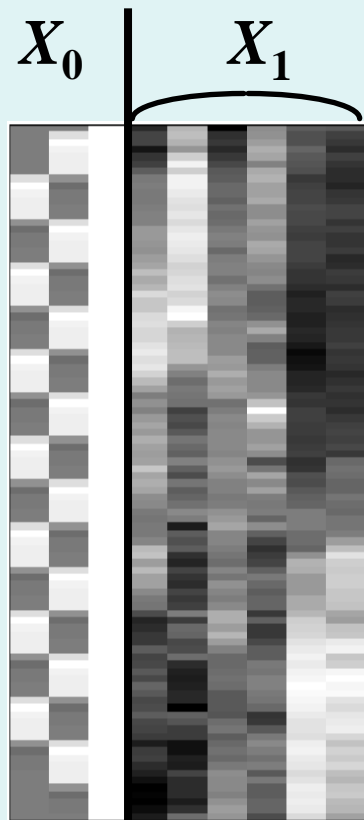
- The  $T$ -statistic does not depend on the scaling of the regressors neither of the contrast.
- Contrast  $c^T \hat{\beta}$  does depend on scaling.
- Be careful of the interpretation of the contrasts  $c^T \hat{\beta}$  themselves (e.g., for a second level analysis):

**sum  $\neq$  average**

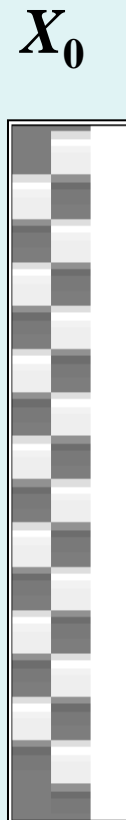
# F-test, extra-sum-of-squares principles

Model comparison:

**Null Hypothesis  $H_0$ :** True model is  $X_0$  (reduced model)



$$\rightarrow \text{RSS} \\ \sum \hat{\varepsilon}_{full}^2$$



$$\rightarrow \text{RSS}_0 \\ \sum \hat{\varepsilon}_{reduced}^2$$

**Test statistic:** ratio of explained variability and unexplained variability (error)

$$F \propto \frac{RSS_0 - RSS}{RSS}$$

$$F \propto \frac{ESS}{RSS} \sim F_{v_1, v_2}$$

$$v_1 = \text{rank}(X) - \text{rank}(X_0) \\ v_2 = N - \text{rank}(X)$$

Full model ?

or Reduced model?



# F-test, multidimensional contrast

Tests *multiple* linear hypotheses:

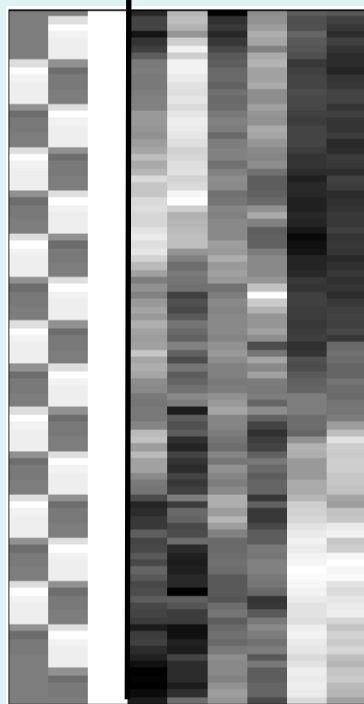
$H_0$ : True model is  $X_0$

$H_0$ :  $\beta_4 = \beta_5 = \dots = \beta_9 = 0$

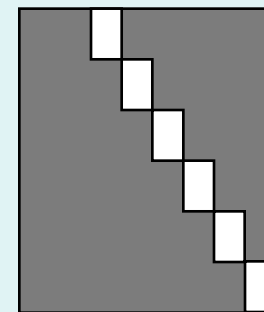
test  $H_0$ :  $c^T \beta = 0$  ?

$X_0$  |  $X_1$  ( $\beta_{4-9}$ )

$X_0$

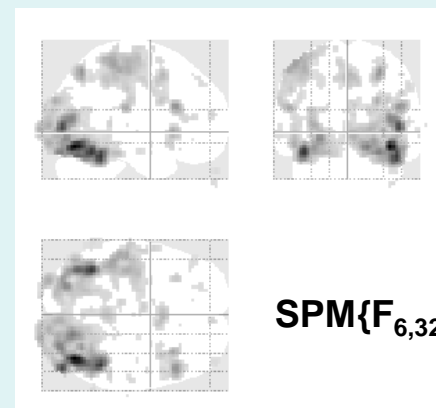
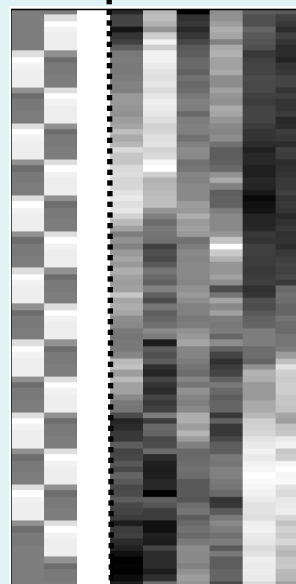


$$c^T = \begin{matrix} 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 \end{matrix}$$



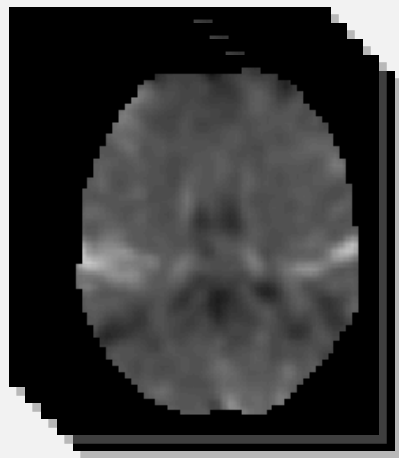
Full model?

Reduced model?



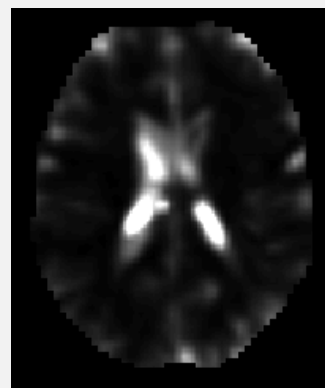
SPM{ $F_{6,322}$ }

# F-contrast in SPM



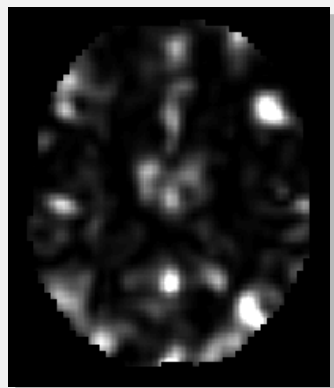
beta\_???? images

$$\hat{\beta} = (X^T X)^{-1} X^T y$$



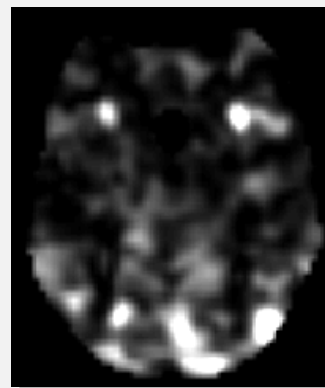
ResMS image

$$\hat{\sigma}^2 = \frac{\hat{\varepsilon}^T \hat{\varepsilon}}{N - p}$$



ess\_???? images

$$(RSS_0 - RSS)$$

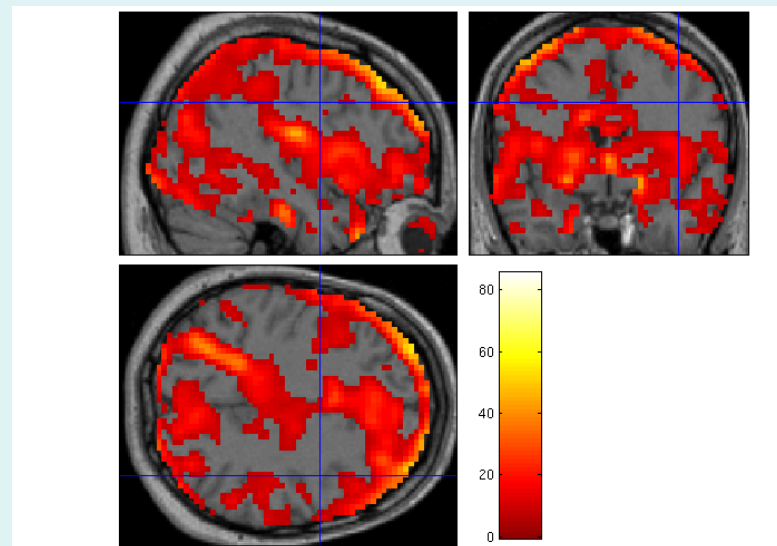
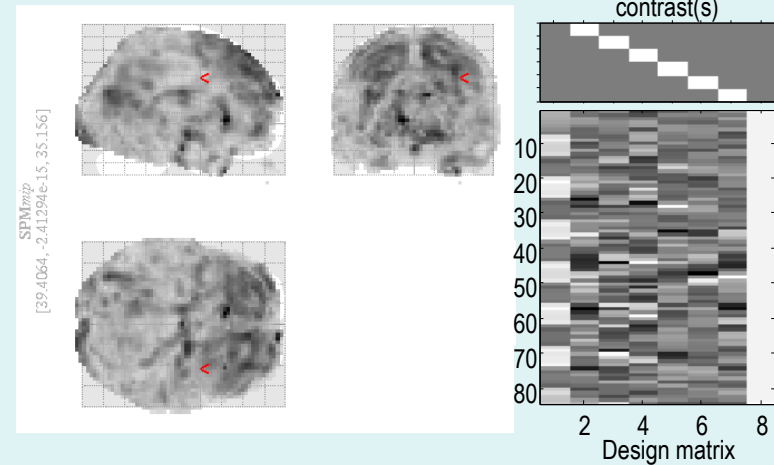
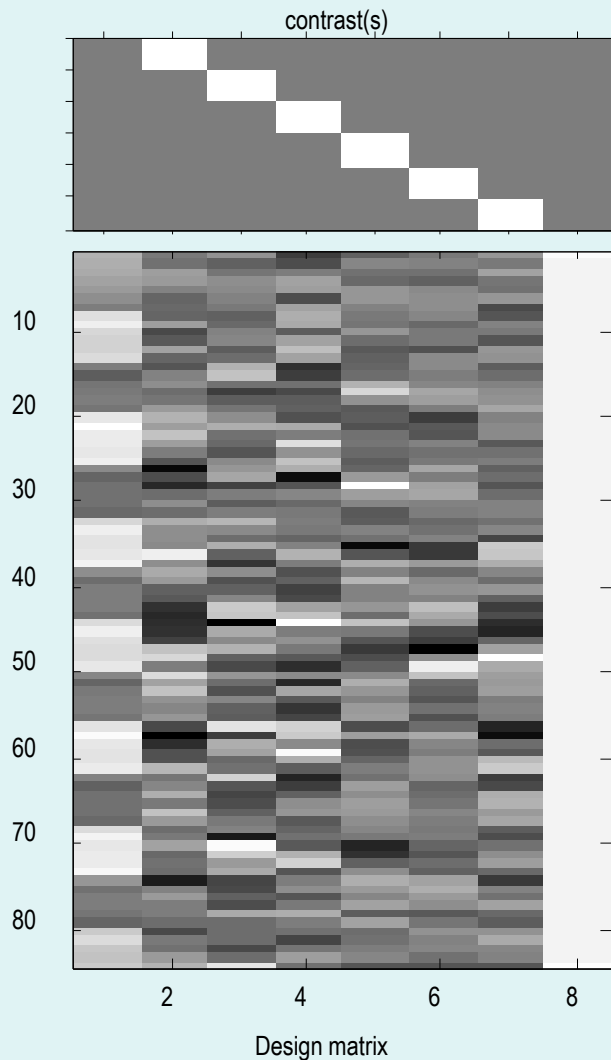


spmF\_???? images

$$SPM\{F\}$$

# F-test, example

## Movement related effects



# F-test summary

- F-tests can be viewed as testing for the additional variance explained by a larger model w.r.t. a simpler (*nested*) model → **model comparison**.
- F-tests a weighted **sum of squares** of one or several combinations of the coefficients  $\beta$ .
- In practice, noneed to explicitly separate X into  $[X_1 X_2]$  thanks to multidimensional contrasts.
- Hypotheses:  
Null Hypothesis  $H_0 : \beta_1 = \beta_2 = \beta_3 = 0$   
Alternative Hypothesis  $H_A : \text{at least one } \beta_k \neq 0$
- In 1D contrast with an F-test, testing  $\beta_1 - \beta_2$  is the same as testing  $\beta_2 - \beta_1$ .

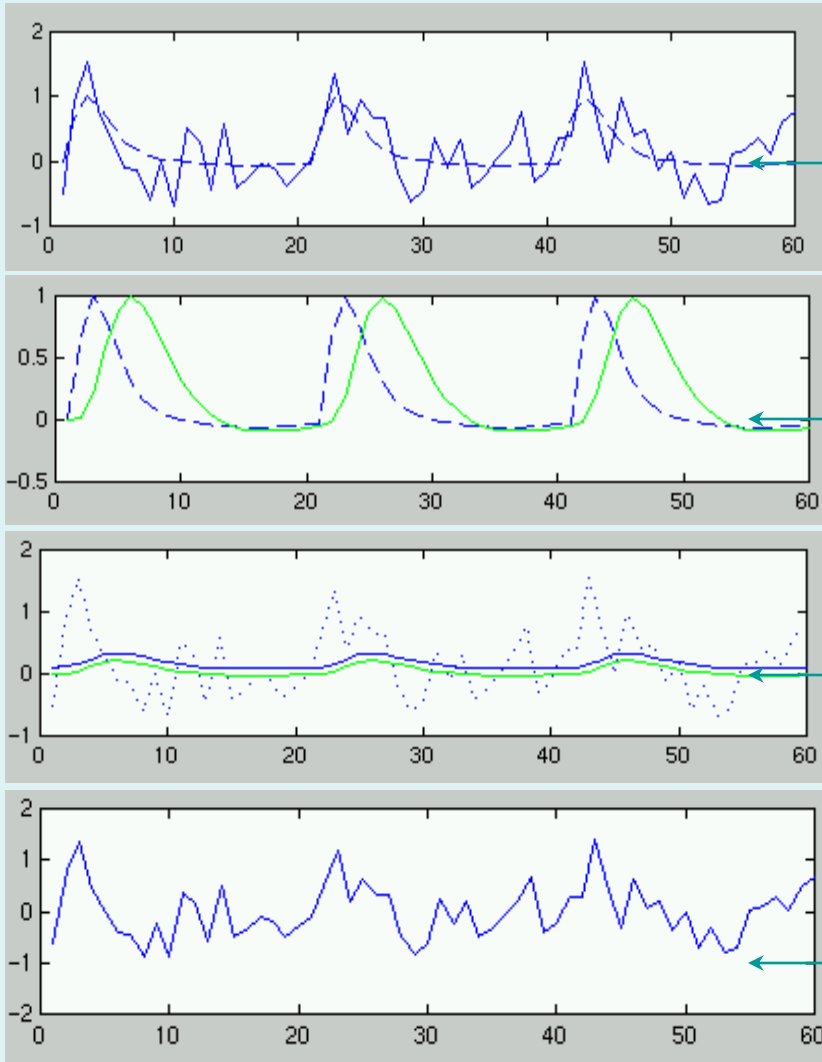
$$\begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 \end{bmatrix}$$

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- **Introduction**
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# A bad model



True signal (---) and observed signal

Model (green, peak at 6sec) and TRUE signal (blue, peak at 3sec)

Fitting :

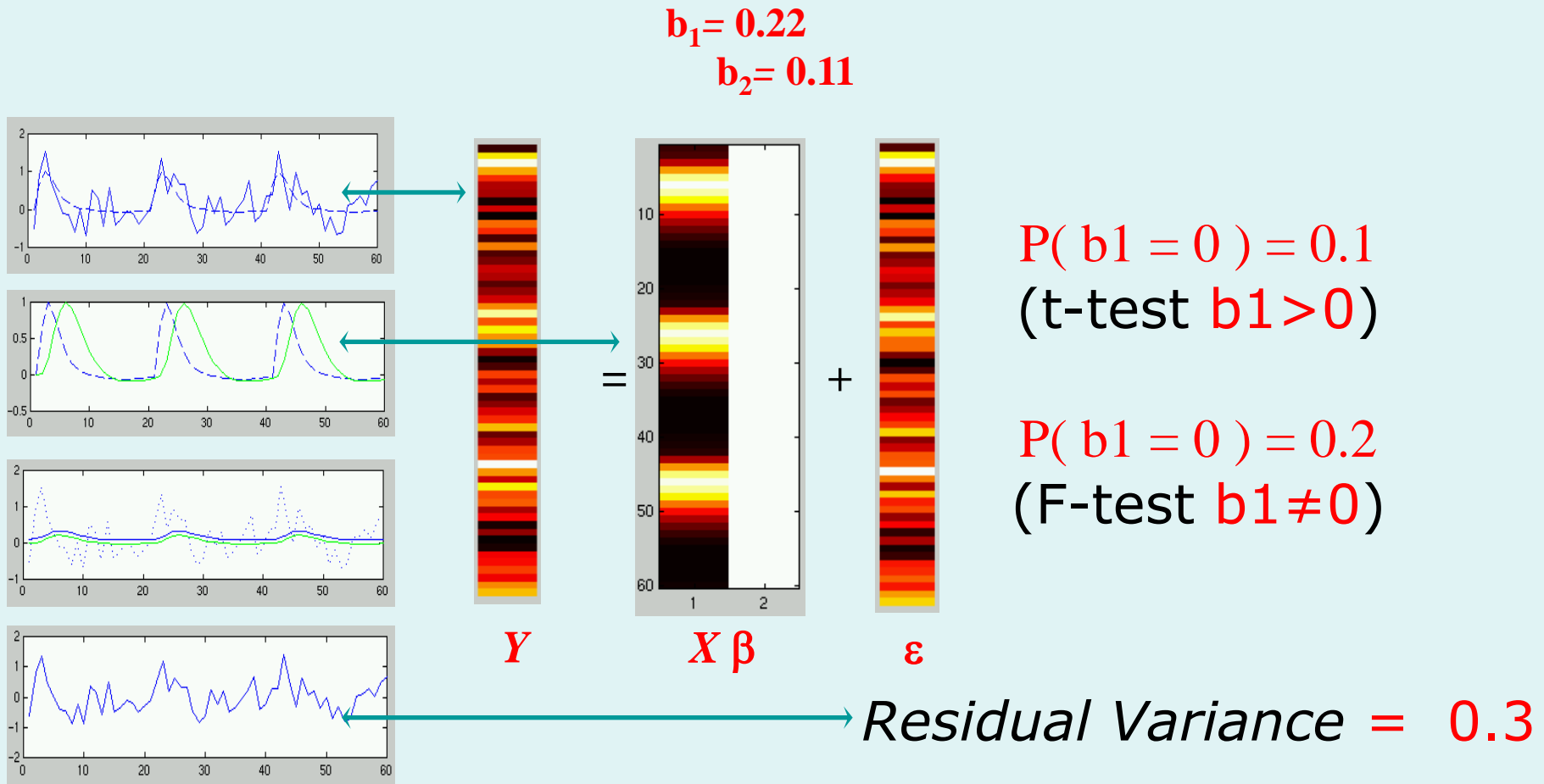
$$b1 = 0.2, \text{ mean} = .11$$

Noise

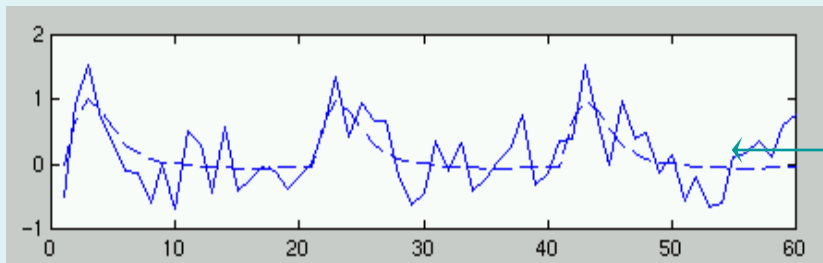
(still contains some signal)

⇒ Test for the green regressor not significant

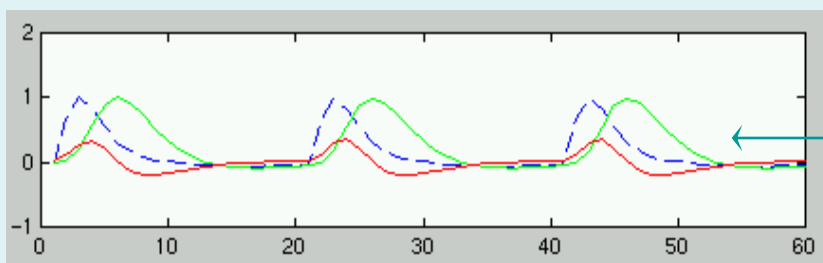
# A bad model



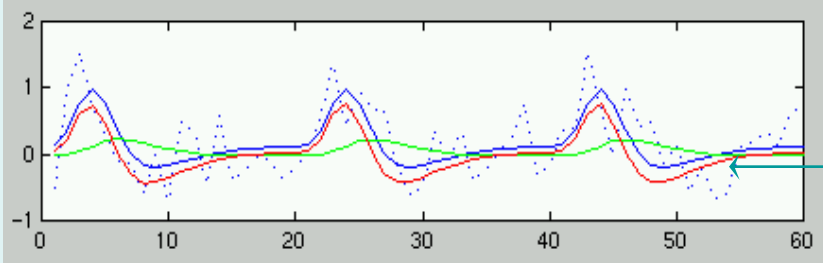
# A better model...



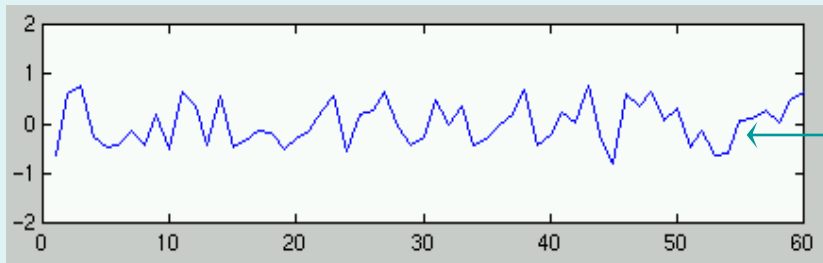
True signal + observed signal



Model (green and red) and true signal (blue ---)  
Red regressor : temporal derivative of the green regressor



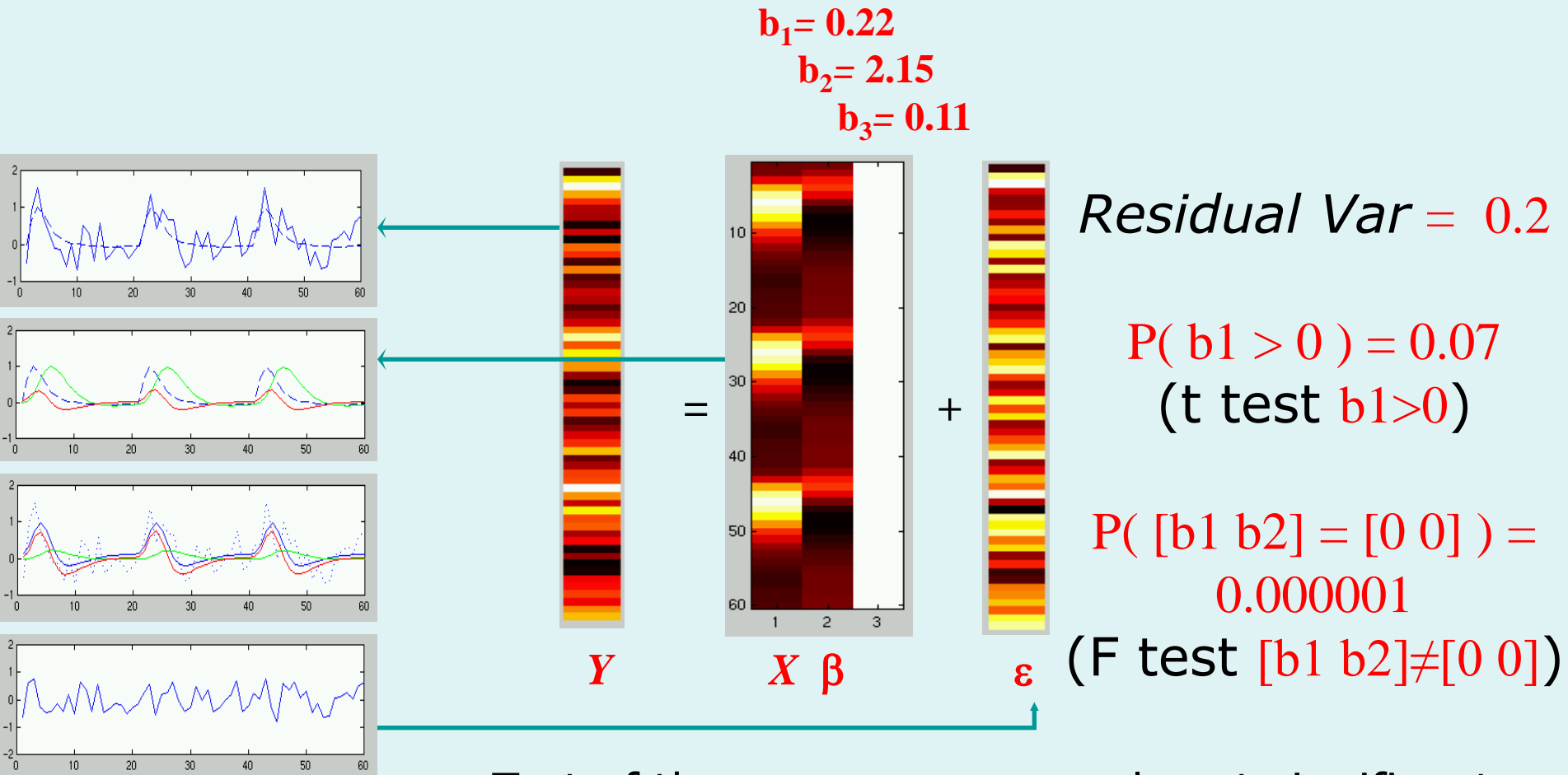
Global fit (blue) and partial fit (green & red)  
Adjusted and fitted signal



Noise (a smaller variance)

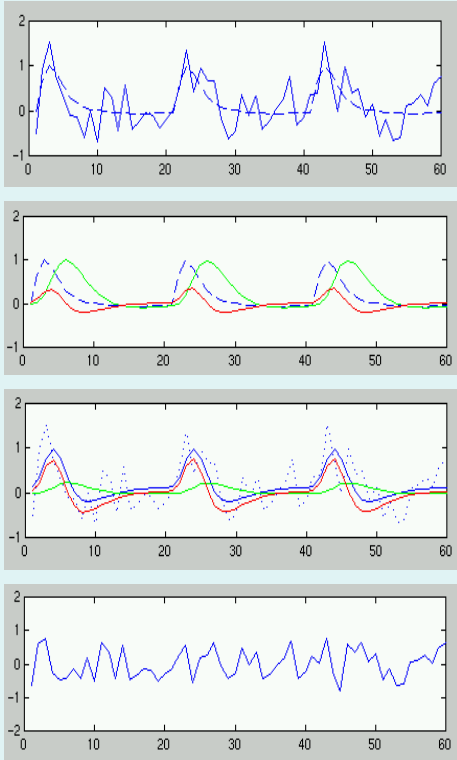


# A better model



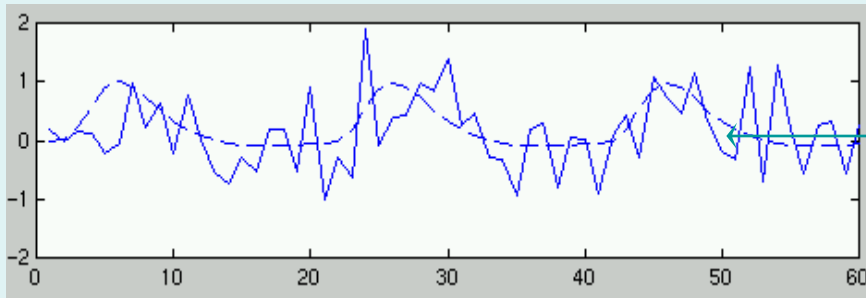
- ⇒ Test of the green regressor almost significant
- ⇒ Test F very significant
- ⇒ Test of the red regressor very significant

# Summary

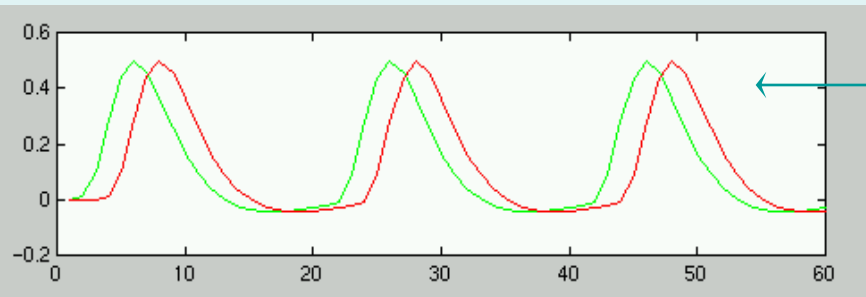


- *The residuals should be looked at ... (non random structure ?)*
- *We rather test flexible models if there is little a priori information, and precise ones with a lot a priori information*
- *In general, use the F-tests to look for an overall effect, then look at the betas or the adjusted signal to characterise the origin of the signal*
- *Interpreting the test on a single parameter (one function) can be very confusing: cf. the delay or magnitude situation*

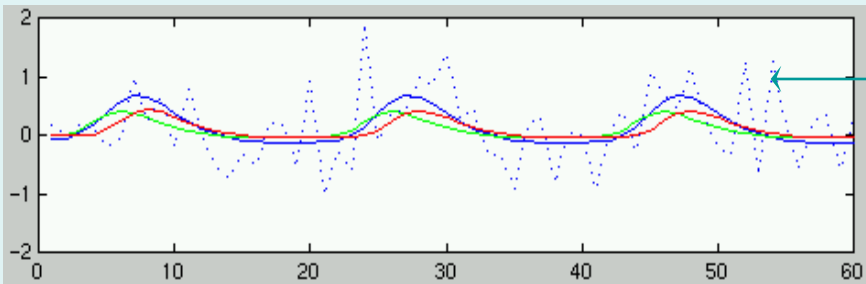
# Correlation between regressors



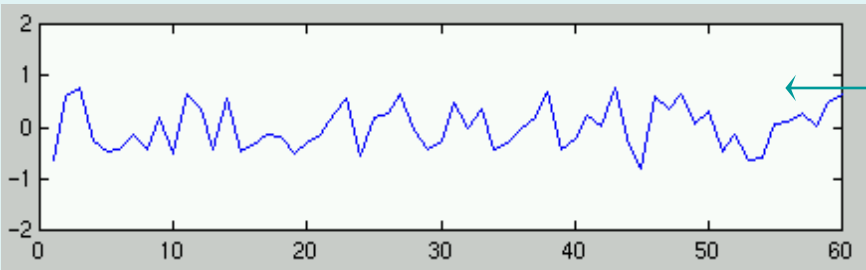
True signal



Model (green and red)

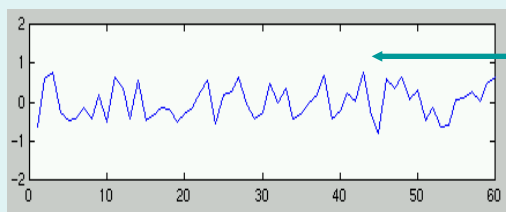
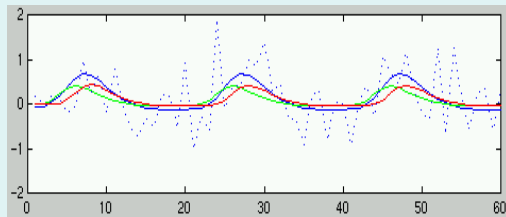
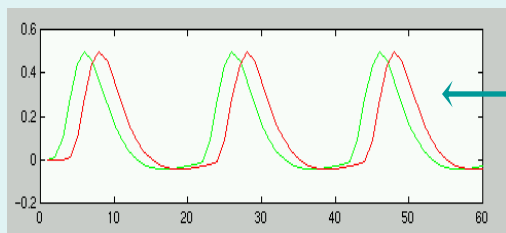
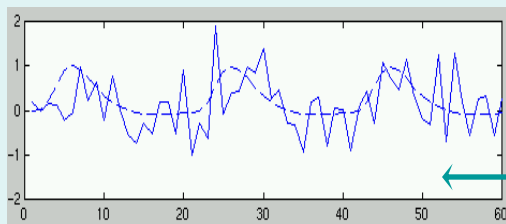


Fitting (blue : global fit)



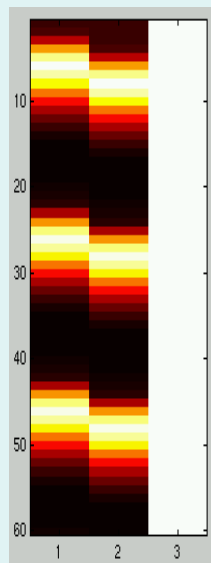
Noise

# Correlation between regressors



$Y$

=



$X \beta$

+



$\epsilon$

$b_1 = 0.79$   
 $b_2 = 0.85$   
 $b_3 = 0.06$

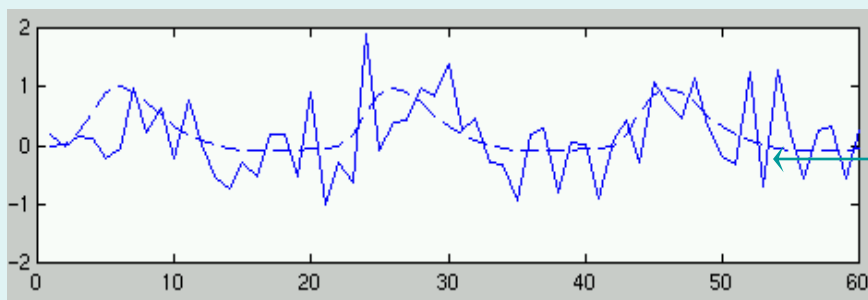
*Residual var.* = 0.2

$P( b_1 = 0 ) = 0.08$   
 (t test  $b_1 > 0$ )

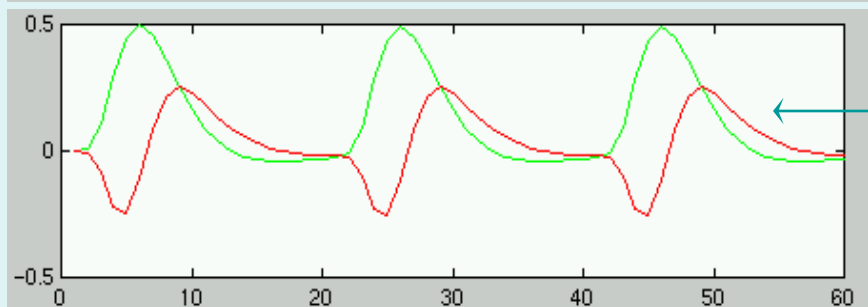
$P( b_2 = 0 ) = 0.07$   
 (t test  $b_2 > 0$ )

$P( [b_1 \ b_2] = 0 ) = 0.002$   
 (F test  $[b_1 \ b_2] \neq 0$ )

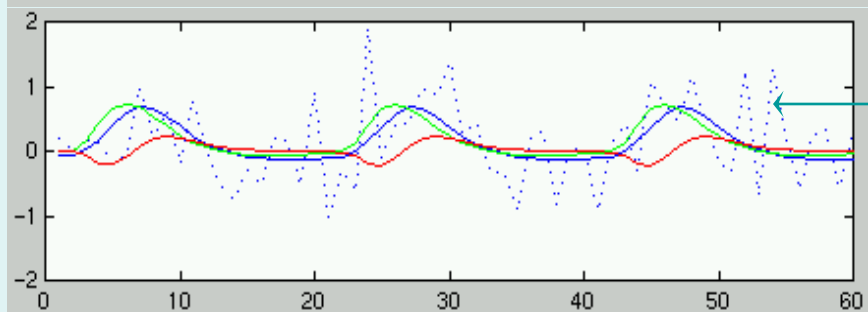
# Decorrelated regressors



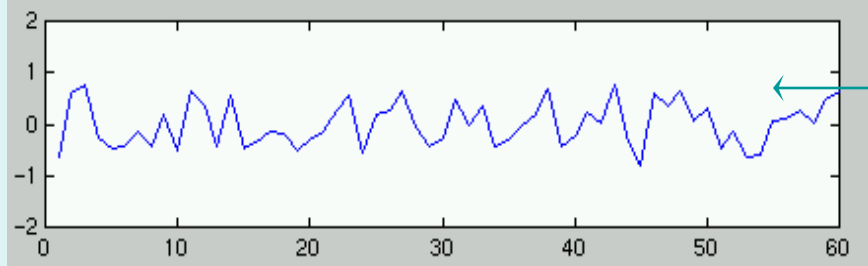
→ true signal



Model : red regressor  
orthogonalised with respect to  
the green one = remove every  
thing that can correlate with  
the green regressor

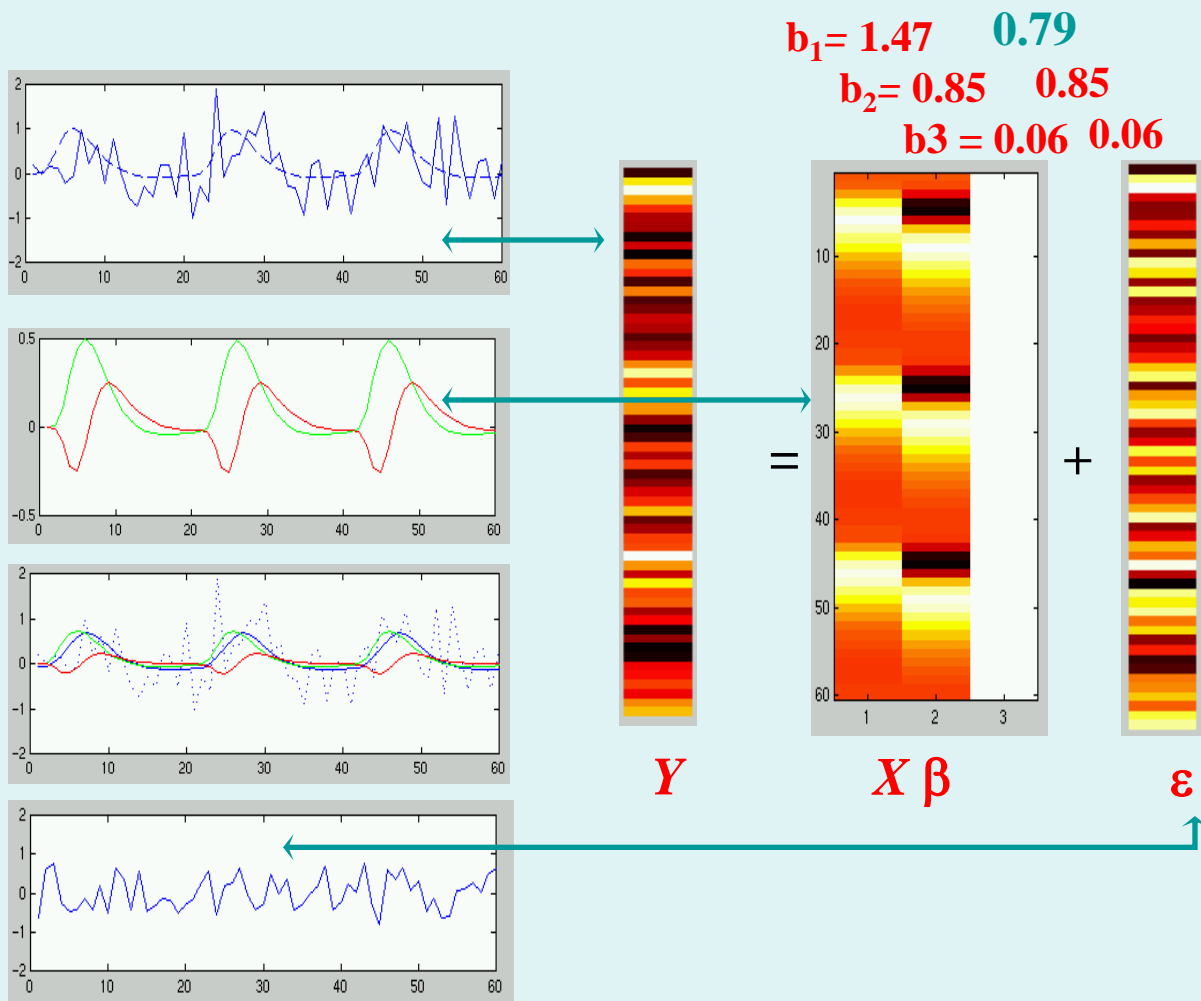


→ Fit



→ Noise

# Decorrelated regressors



*Residual var. = 0.2*

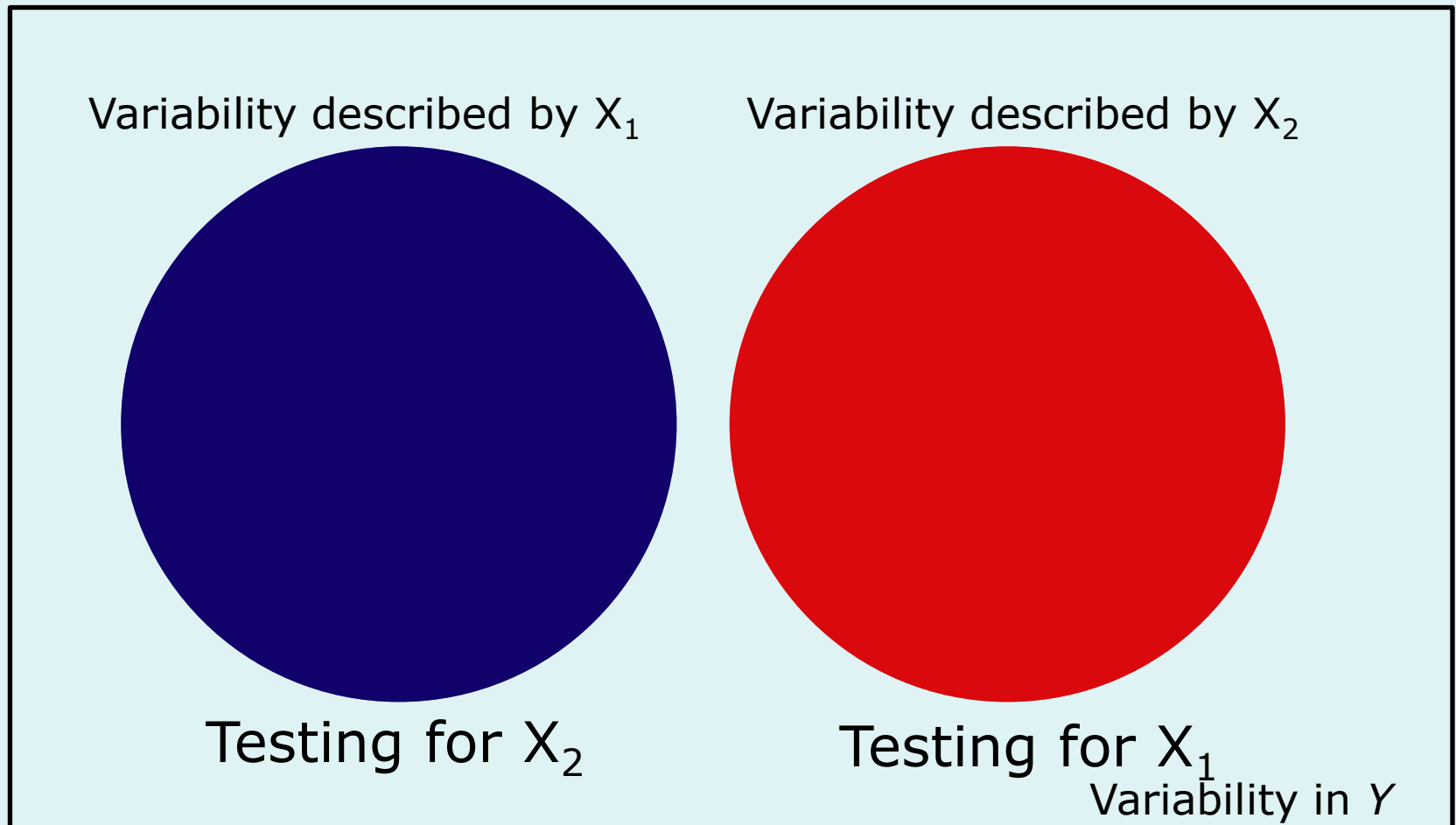
$P(b_1 = 0) = 0.0003$   
(t test  $b_1 > 0$ )

$P(b_2 = 0) = 0.07$   
(t test  $b_2 > 0$ )

$P([b_1 \ b_2] = 0) = 0.002$   
(F test  $[b_1 \ b_2] \neq 0$ )

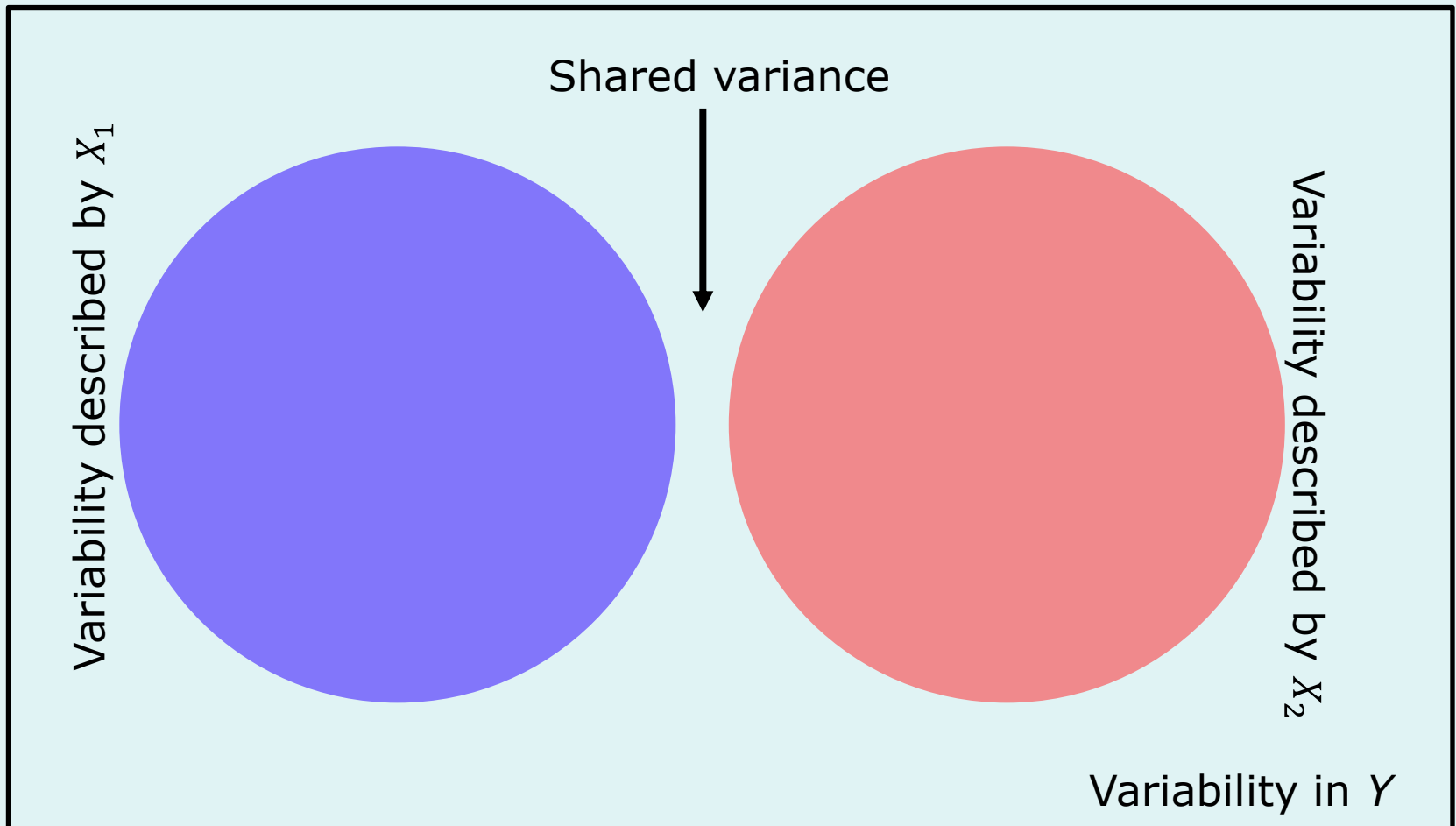
# Orthogonal regressors

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# Orthogonal regressors

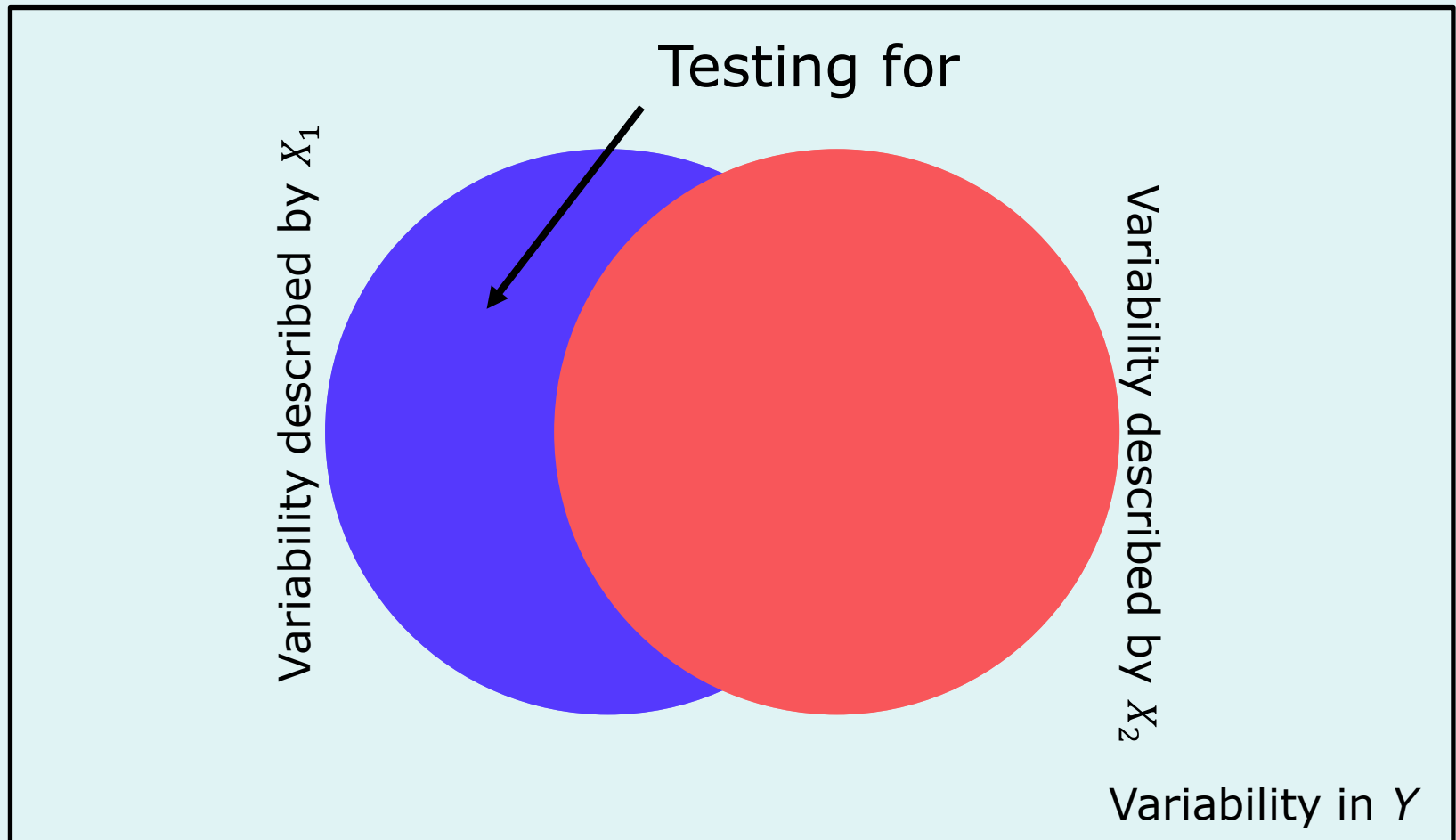
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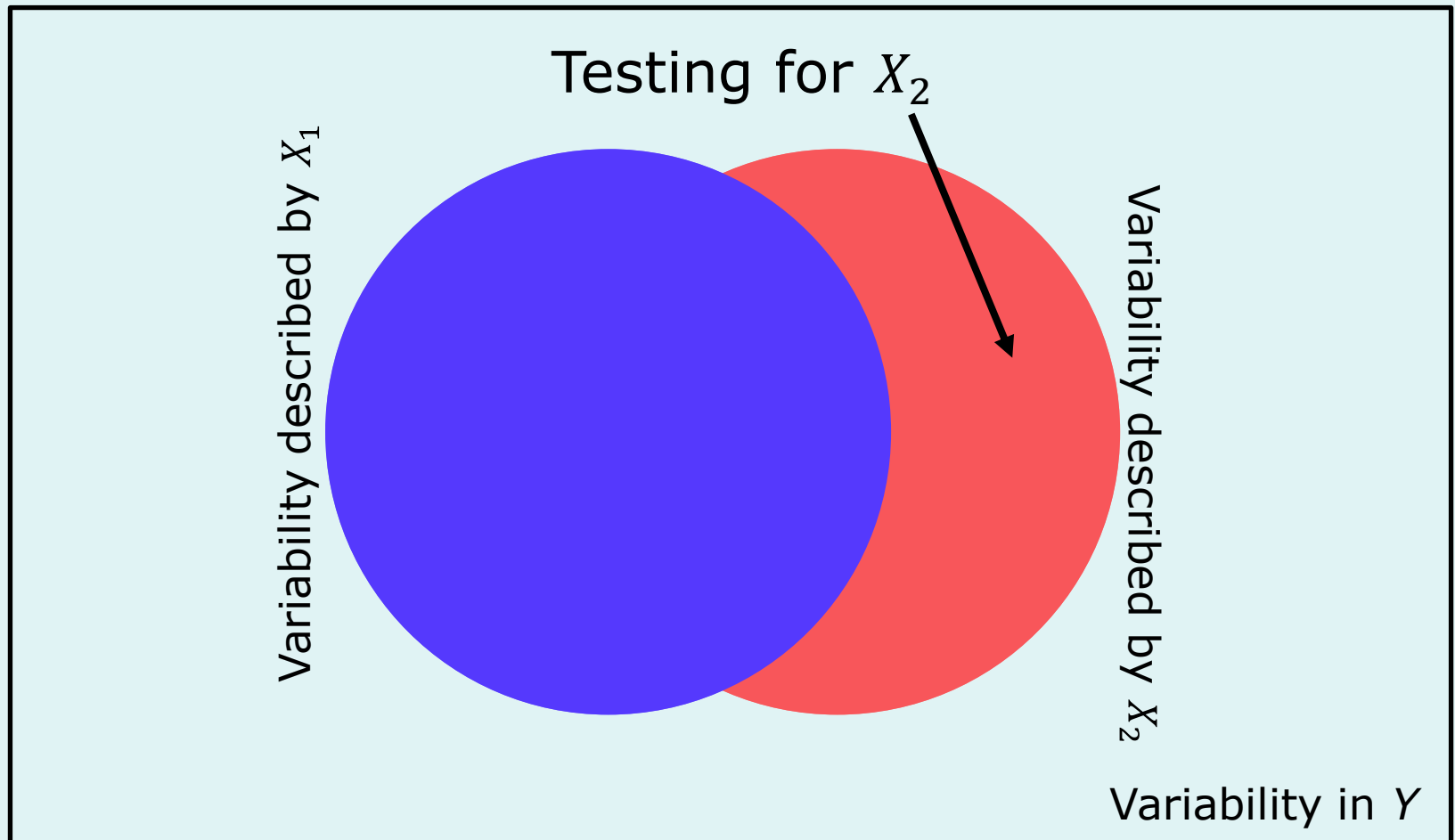
# Orthogonal regressors

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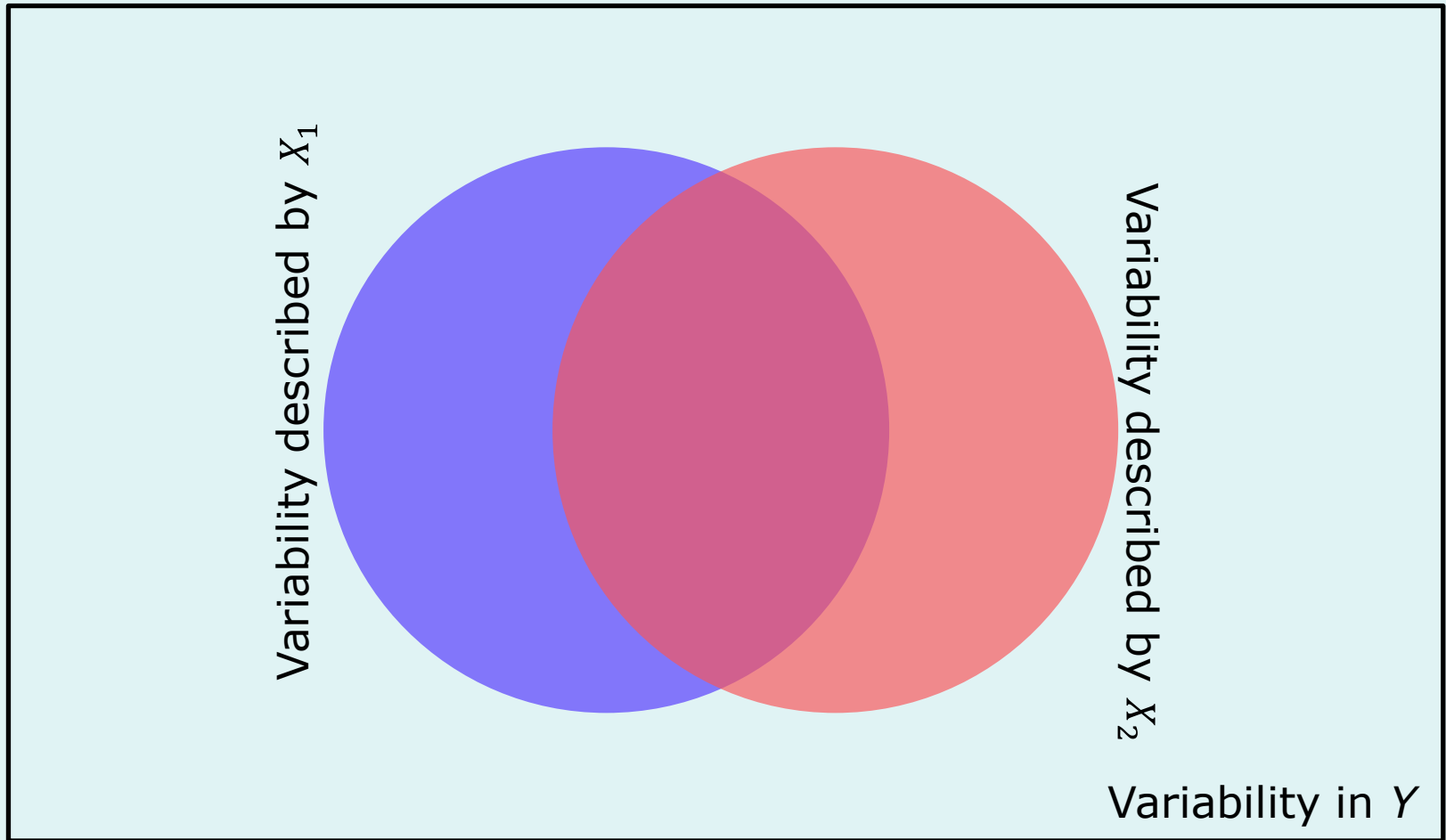
# Orthogonal regressors

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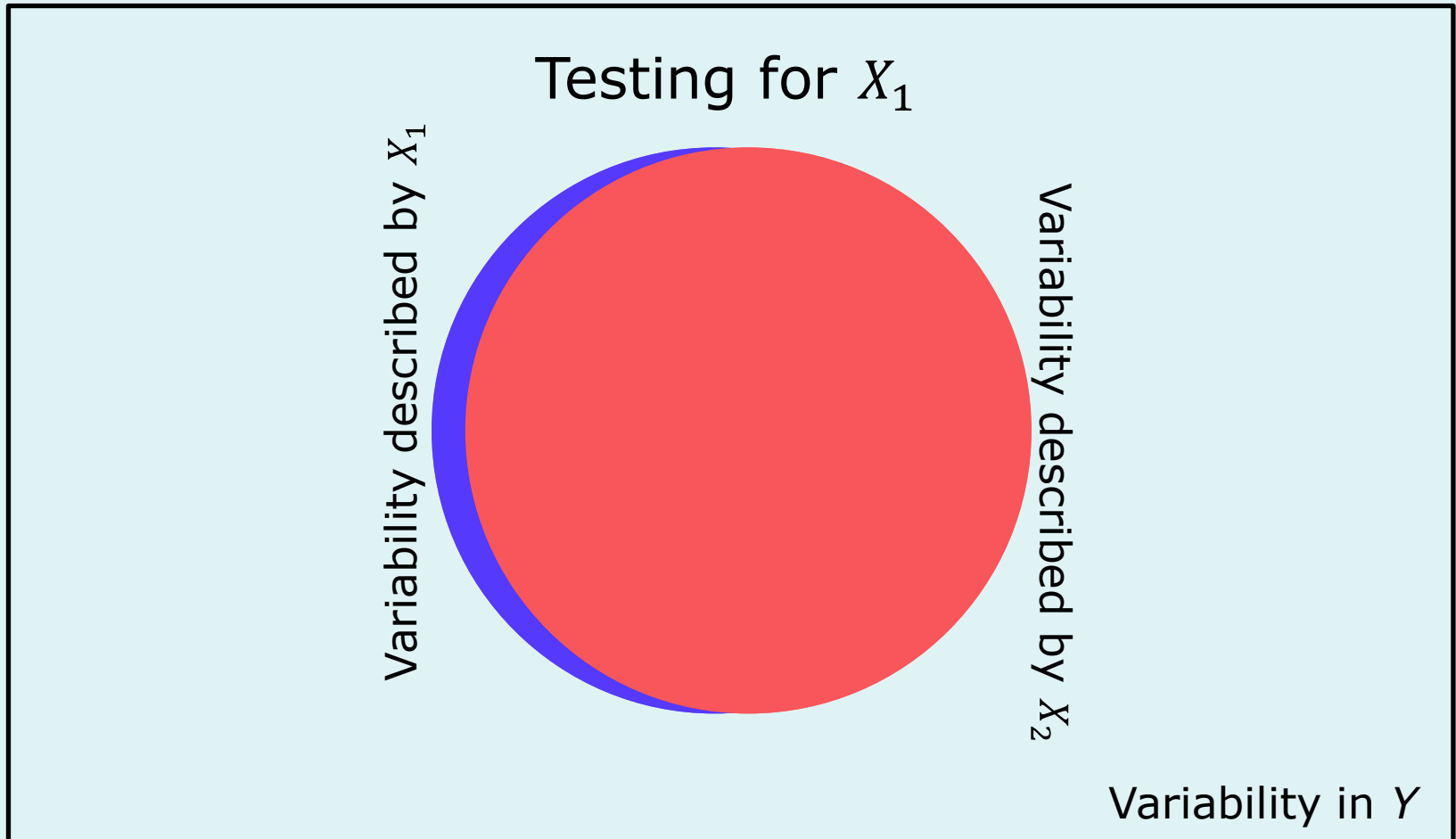
# Orthogonal regressors

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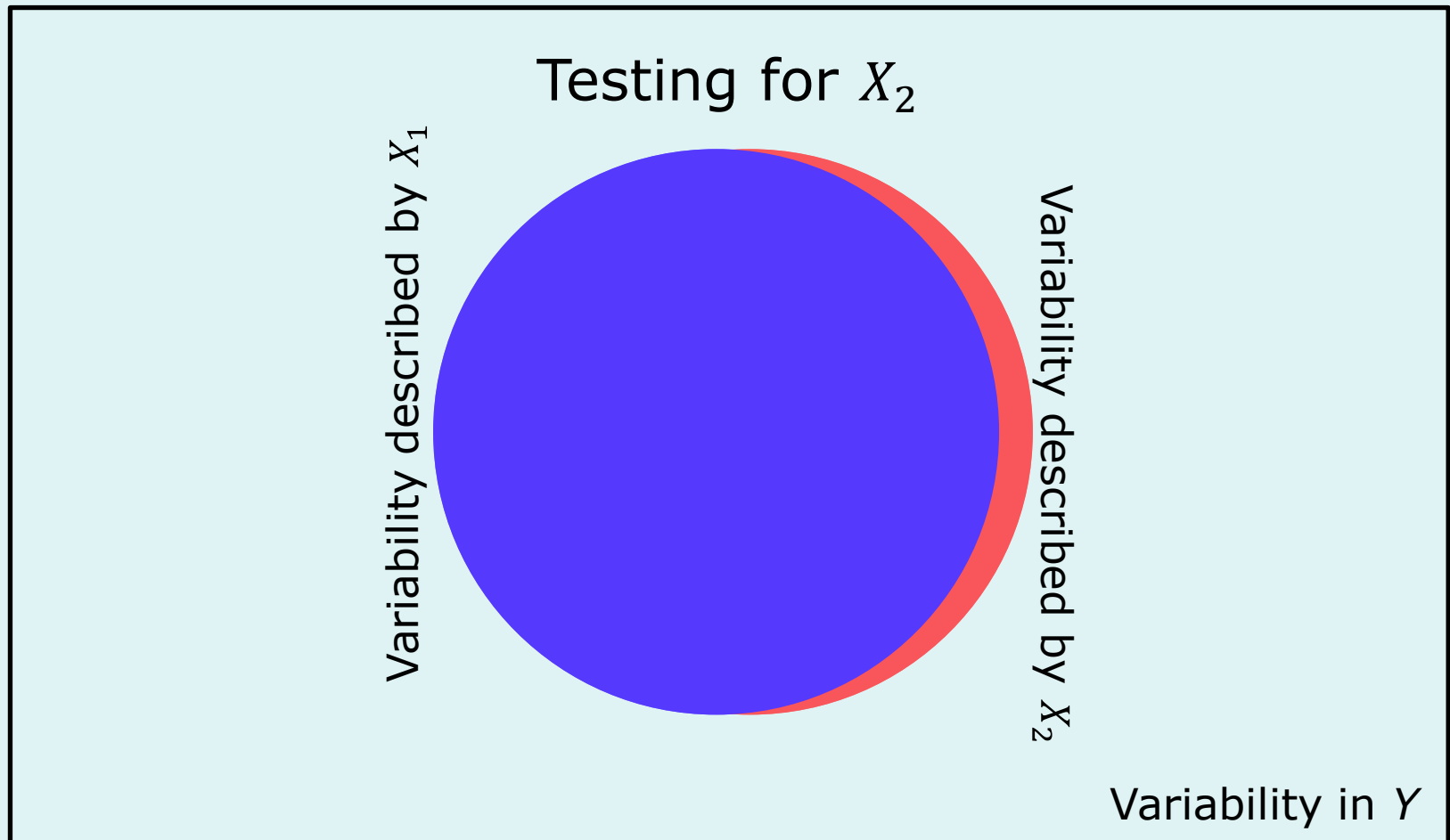
# Orthogonal regressors

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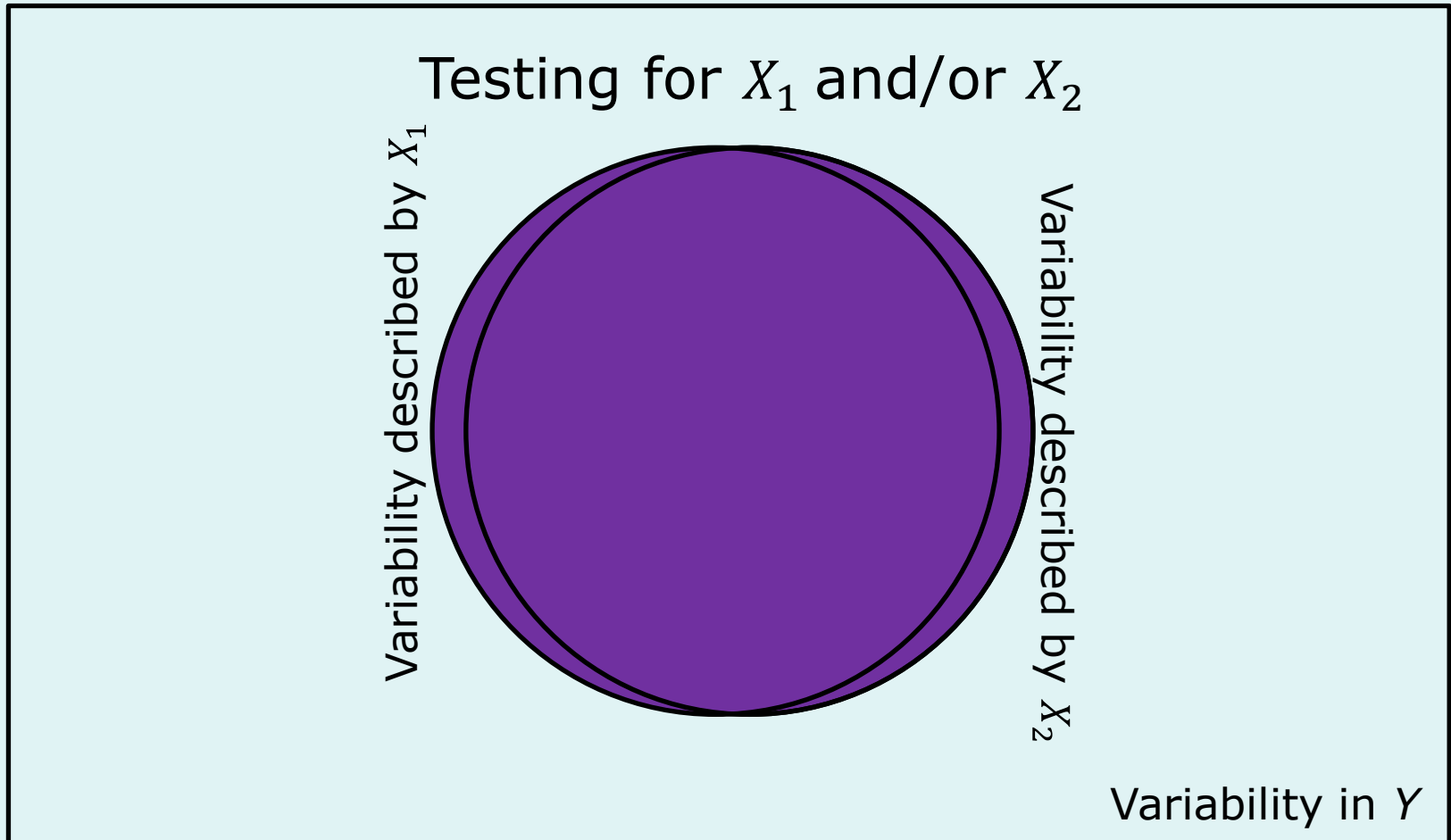
# Orthogonal regressors

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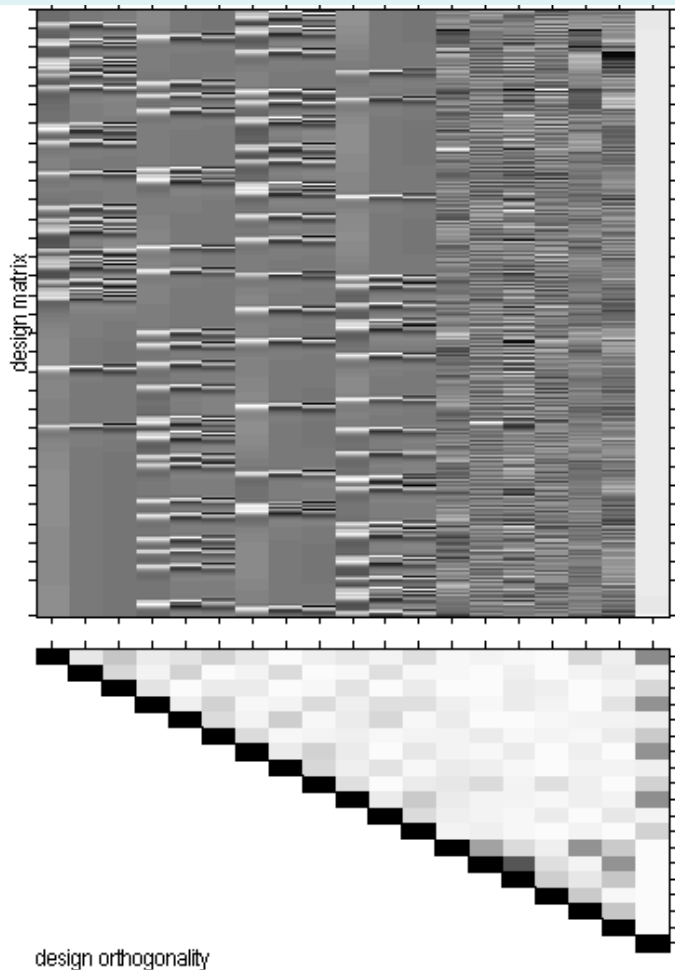


# Orthogonal regressors

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# Design orthogonality

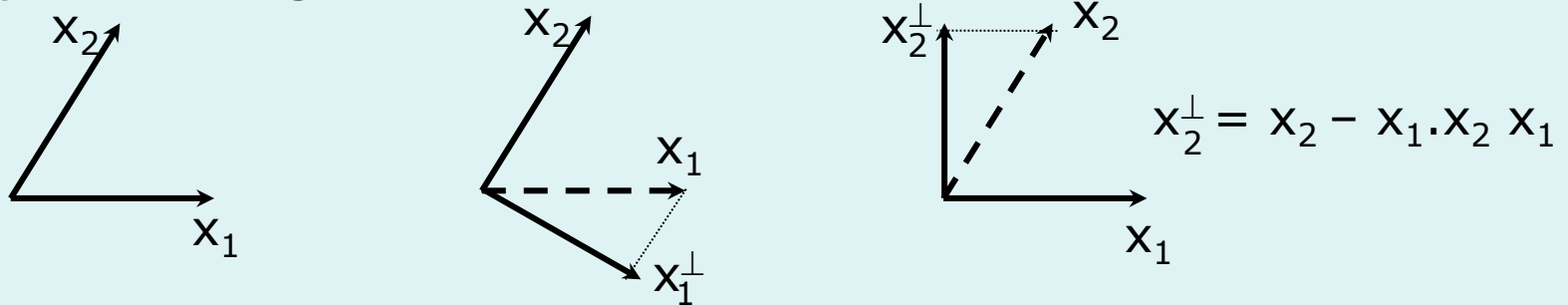


**Measure** : abs. value of cosine of angle between columns of design matrix  
**Scale** : black - colinear ( $\cos=+1/-1$ )  
white - orthogonal ( $\cos=0$ )  
gray - not orthogonal or colinear

- For each pair of columns of the design matrix, the orthogonality matrix depicts the magnitude of the **cosine of the angle** between them, with the range 0 to 1 mapped from white to black.
- If both vectors have **zero mean** then the cosine of the angle between the vectors is the same as the **correlation** between the two variates.

# Correlated regressors

- We implicitly test for an additional effect only. When testing for the first regressor, we are effectively removing the part of the signal that can be accounted for by the second regressor → implicit orthogonalisation.



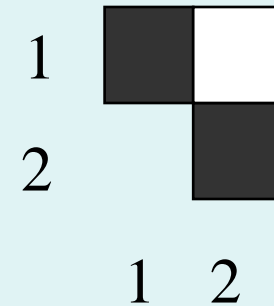
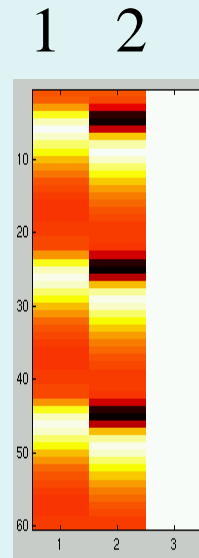
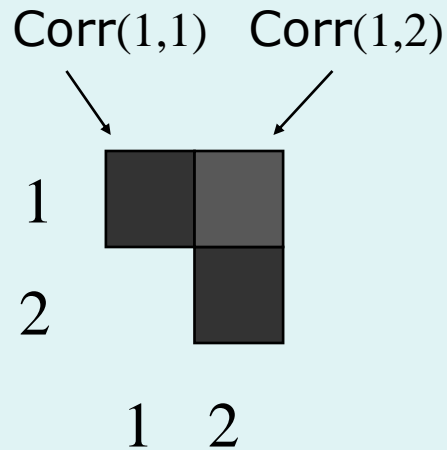
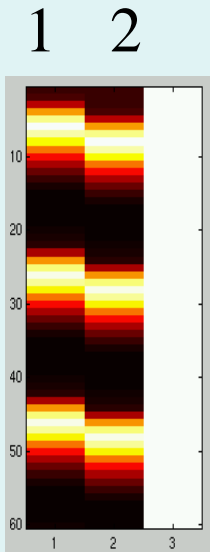
- Orthogonalisation = decorrelation. Parameters and test on the non modified regressor change.  
Rarely solves the problem as it requires assumptions about which regressor to uniquely attribute the common variance.  
→ change regressors (i.e. design) instead, e.g. factorial designs.  
→ use F-tests to assess overall significance.
- Original regressors may not matter: it's the contrast you are testing which should be as decorrelated as possible from the rest of the design matrix



# Design orthogonality

Black = completely correlated

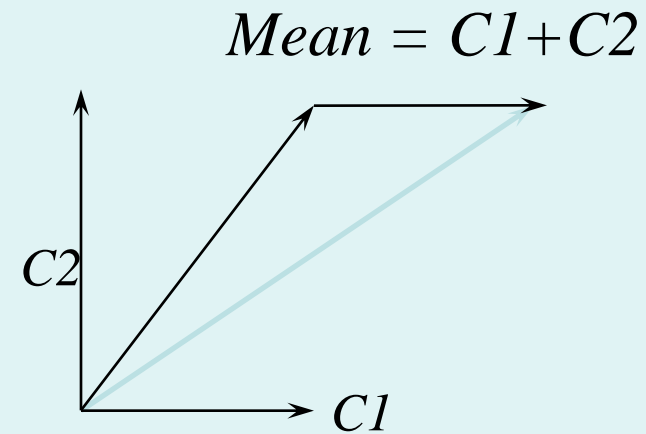
White = completely orthogonal



*Beware* : when there is more than 2 regressors ( $C_1, C_2, C_3, \dots$ ), you may think that there is little correlation (light grey) between them, but  $C_1 + C_2 + C_3$  may be correlated with  $C_4 + C_5$

# Rank-deficient model

$$Y = Xb + e$$
$$X = \begin{matrix} & 1 & 0 & 1 \\ & 0 & 1 & 1 \\ & 1 & 0 & 1 \\ & 0 & 1 & 1 \\ \begin{matrix} \nearrow & \uparrow & \nwarrow \\ \text{Cond 1} & \text{Cond 2} & \text{Mean} \end{matrix} \end{matrix}$$



Parameters are **not unique** in general !

Some contrasts have no meaning: **NON ESTIMABLE**

Example here :

- $c' = [1 \ 0 \ 0]$  is not estimable  
( = no specific information in the first regressor);
- $c' = [1 \ -1 \ 0]$  is estimable.

# Summary

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- We are implicitly testing additional effect only, so we may miss the signal if there is some correlation in the model using t tests
- Orthogonalisation is not generally needed - parameters and test on the changed regressor don't change
- It is always simpler (when possible !) to have orthogonal (uncorrelated) regressors
- In case of correlation, use F-tests to see the overall significance. There is generally no way to decide where the « common » part shared by two regressors should be attributed to
- In case of correlation and you need to orthogonalise a part of the design matrix, there is no need to re-fit a new model : the contrast only should change.

# Content

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- **Introduction**
- **Contrast & Inference**
- **Orthogonality issue**
- **Conclusion**

# Way to proceed

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Prepare your questions.

**ALL the questions !**



Find a model which

- allows contrasts that translates these questions.
- takes into account ALL the effects (interaction, sessions, etc)



Devise task & stimulus presentation.



Acquire the data & analyse.



**Not the other way round!!!**

# References

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- Statistical Parametric Mapping: The Analysis of Functional Brain Images. Elsevier, 2007.
- Plane Answers to Complex Questions: The Theory of Linear Models. R. Christensen, Springer, 1996.
- Statistical parametric maps in functional imaging: a general linear approach. K.J. Friston et al, Human Brain Mapping, 1995.
- Ambiguous results in functional neuroimaging data analysis due to covariate correlation. A. Andrade et al., NeuroImage, 1999.

